

Revisiting the Robustness of (R/M)LWR under Polynomial Moduli with Applications to Lattice-Based Compact SO-CCA Security

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Abstract. This work conducts a comprehensive investigation on determining the *entropic* hardness of (R/M)LWR under polynomial modulus. Particularly, we establish the hardness of (M)LWR for general entropic secret distributions from (Module) LWE assumptions based on a new conceptually simple framework called rounding lossiness. By combining this hardness result and a trapdoor inversion algorithm with asymptotically the most compact parameters, we obtain a compact lossy trapdoor function (LTF) with improved efficiency. Extending our LTF with other techniques, we can derive a compact all-but-many LTF and PKE scheme against selective opening and chosen ciphertext attacks, solely based on (Module) LWE assumptions within a polynomial modulus. Additionally, we show a search-to-decision reduction for RLWR with Gaussian secrets from a new Rényi Divergence-based analysis.

1 Introduction

Lattice-based cryptography has attracted significant attention in recent years – first it stands out as one of very few promising candidates against quantum algorithms [52], and moreover, it provides as a robust foundation upon which a wide array of (advanced) crypto systems can be built, e.g., [44]. Particularly, many lattice-based crypto systems are directly based on the *learning with error* (LWE) problem [49], which enjoys search-to-decision reductions [37, 38, 41, 49] and as well as worst-case hardness from some lattice problems, under quantum or classical reductions [15, 41, 49]. These results instill confidence in the hardness of LWE, encompassing both its decision and search forms, and consequently, in the security of cryptographic systems derived from LWE.

However the LWE problem requires to sampling random errors, leading to efficiency losses and complications in designing some cryptographic primitives that are deterministic in its nature of computation, e.g., pseudorandom functions (PRFs). To tackle these challenges, the work [5] introduced the Learning with Rounding (LWR) problem as a derandomized version of the LWE. Then the research community identified that many crypto systems can be naturally

Compact LTFs from Lattices. Lossy trapdoor functions (LTFs) are powerful crypto tools that can be used to construct many applications, such as trapdoor one-way functions, collision-resistant hash functions, lossy encryption, CCA2 secure PKE, etc, [47]. They can be extended to design the more advanced *all-but-many lossy trapdoor functions* (ABM-LTFs) [12, 32], which can be used to realize PKE with a stronger notion of security, namely, *selective opening chosen-ciphertext-security* (SO-CCA security). For lattice-based constructions, there are several prior works [1, 7, 21, 47], among which, the construction in [1] is conceptually very simple, based on entropic LWR.

However, all the prior schemes have some drawbacks across various aspects, including *information rate*⁴, *lossiness*⁵, and the *public parameter size*. For example, the constructions in [1, 47] suffer from super-constant information rate, the work [7] achieves small lossiness parameter, and the work [21] requires very large public parameters and involves much complicated parallel repetition, leading to poor efficiency.

Recently, the work [27] designed a compact LTF ($O(1)$ information rate) based on lattice, and further extended the basic result to compact *all-but-many*-LTFs (ABM-LTFs) and *selective opening*-CCA (SO-CCA) secure PKE based. Despite the theoretical advancements, the designs in [27] are rather intricate, involving heavy Gaussian sampling, harsh restriction for achieving strong lossiness, and large public parameters compared to the very simple (though non-compact) [1]. Additionally, their ABM-LTF construction assumes the existence of a PRF computable in NC1 (taking the PRF key as input). While such a PRF can be instantiated from lattice [11], the construction requires a super-polynomial modulus. Consequently, it remains unclear whether the results of [27] can be derived from lattices under a polynomial modulus, which is a weaker assumption. These considerations motivates our second goal.

(**Main Goal 2:**) Improve the state of the art of LTF constructions in [27] with a more conceptually straightforward and efficient design. Then determine whether we can eliminate the requirement of PRF in NC1 to achieve ABM-LTF and SO-CCA PKE under a polynomial modulus.

1.1 Our Contributions

This work aims at the two main goals and makes three major contributions.

Contribution 1. We establish hardness results for the general entropic LWR problem from the standard lattice-based assumption. In particular, we show a reduction from LWE to entropic LWR with general entropic secret distributions (i.e., which only require sufficient entropy over the secret). To achieve this, we

⁴ Information rate is defined as the input-to-output ratio, with a higher value being preferable. A design is *compact* if the rate is $O(1)$.

⁵ Lossiness is the parameter that quantifies the average number of bits lost when evaluating the function in the lossy mode. In our applications, a higher lossiness is desirable.

Compared with the constructions in [27], our constructions have several significant advantages. Firstly, our basis LTF scheme is much simpler, due to the extremely simple evaluation algorithms of LWR. Secondly, the amount of “lossiness” in our LTF construction is more flexible. Particularly, we can achieve the *relative lossiness*⁷ arbitrarily close to 1 with poly modulus rather than super-poly modulus in [27]. Next, our LTF is with smaller evaluation key, i.e., $O(n^2 \log q)$ for ours vs $O(n^2 \log^4 q)$ for [27]. Finally, our constructions of compact ABM-LTFs and compact SO-CCA secure PKE are with polynomial modulus without relying on the additional assumption of PRFs mentioned in [27].

Contribution 3. We prove pseudorandomness of RLWR with Gaussian secret from the standard assumptions over ideal lattices. Particularly, we first show a reduction from search RLWR with certain entropic secret distributions (with sufficient entropy) to decision RLWR with Gaussian secret distributions. To the best of our knowledge, this is the first hardness result that captures RLWR with bounded secret distributions. The crux is a Rényi Divergence (RD)-based noise flooding technique for matrix-vector multiplication (or multiplication of ring elements). As previous analyses mainly focus on the vector addition (or ring elements addition), our new analysis would be of independent interest. Informally, this hardness result can be summarized as following:

Theorem 1.2 *Assume one-way hardness of RLWR with certain entropic secret distribution holds under poly-modulus (and for other appropriate parameters), then the decision RLWR with Gaussian secret distributions (defined according to coefficient embeddings) also holds.*

We next generalize the search RLWE to search RLWR reduction in [33] to the case of entropic secrets, and thus establish the hardness of our special search entropic RLWR. Combining with the hardness results of entropic RLWE in [14], the hardness of our special search entropic RLWR can be further established from the standard assumptions (e.g. RLWE and NTRU). Informally, we have the following corollary:

Corollary 1.3 *Assume the pseudorandomness of RLWE and NTRU holds under poly-modulus (and for other appropriate parameters), then the decision RLWR with Gaussian secret distributions (defined according to coefficient embeddings) is $\frac{1}{\text{poly}(\lambda)}$ -secure.*

It needs to point out that we cannot bridge the (strong) pseudorandomness (i.e. $\text{negl}(\lambda)$ -security) of RLWR with Gaussian secret distributions to the standard assumptions under poly-modulus. The main technique barrier is that there exists a lower bound for showing a sample-preserving reduction from RLWE to RLWR with polynomial modulus by a recent work [40]. Nevertheless, as mentioned in [33], this barrier for applications can be overcome via the hardness amplification technique of [55].

⁷ The measure of *relative lossiness* in [27] is to denote the ratio of the remaining entropy of the input and the original entropy of the input.

construction and our previous entropic LWR reduction. As a crux analysis in our compact trapdoor design, we first show a theorem about the requirements of a random matrix having an invertible sub-matrix.

5.1 Probability of Matrix with Generalized Invertibility

First of all, we define what it means for a matrix from $\mathbb{Z}_q^{n \times m}$ ($m \geq n$) to be *invertible* for any modulus q . This notion is a generalization of the square matrix's invertibility to a more general case of matrix from $\mathbb{Z}_q^{n \times m}$ ($m \geq n$), and also appeared in several previous works [2, 13].

Definition 5.1 Let $m' \geq n \geq 1$ be dimension parameters and $q \geq 2$ be any modulus. For a matrix $\mathbf{A} = (\mathbf{a}_i)_{i \in [m']} \in \mathbb{Z}_q^{n \times m'}$, we define that a matrix $\mathbf{A} \in \mathbb{Z}_q^{n \times m'}$ is invertible if there exists n positive indexes $i_1 < i_2 < \dots < i_n \leq m'$ such that $\mathbf{H} = (\mathbf{a}_{i_1} | \mathbf{a}_{i_2} | \dots | \mathbf{a}_{i_n}) \in \mathbb{Z}_q^{n \times n}$ is invertible. We define $\varepsilon_{\text{non-inv}}^{n, m', q}$ to be the probability of a uniformly random matrix in $\mathbb{Z}_q^{n \times m'}$ to be not invertible, i.e., $\varepsilon_{\text{non-inv}}^{n, m', q} = \Pr_{\mathbf{A} \leftarrow \mathbb{Z}_q^{n \times m'}}[\mathbf{A} \text{ is not invertible}]$.

Afterwards, we prove that if m' is slightly larger than n , then the probability $\varepsilon_{\text{non-inv}}^{n, m', q}$ is negligible for arbitrary polynomial-size modulus q .

Theorem 5.2 For any modulus $q \geq 25$, $n \geq 1$ and $t_1, t_2 \geq 1$, for $m' = n + t_1 + t_2 \ln \ln q$, we have $\varepsilon_{\text{non-inv}}^{n, m', q} \leq 2^{-(t_1+1)} + e^{-t_2/4}$.

Proof. The general idea of this proof is to separate m' uniform and independent vectors from \mathbb{Z}_q^n into former $(n + t_1)$ vectors and latter $t_2 \ln \ln q$ vectors. We then illustrate that 1. there exists $n - 1$ linearly independent vectors $\{\mathbf{u}_1, \dots, \mathbf{u}_{n-1}\}$ in the first $n + t_1$ vectors except with probability 2^{-t_1} ; 2. given $n - 1$ existing linearly independent vectors, there exists a vector \mathbf{u}_n in the last $t_2 \ln \ln q$ vectors such that \mathbf{u}_n is linearly independent from $\{\mathbf{u}_1, \dots, \mathbf{u}_{n-1}\}$ except with probability $e^{-t_2/4}$.

Let $q = q_1 q_2 \dots q_k$ represents the prime factorization of the modulus q where each $q_j = p_j^{d_j}$ is a power of prime. Let $\{\mathbf{u}_i\}_{1 \leq i \leq n}$ be vectors from \mathbb{Z}_q^n . For $1 \leq i \leq n$, denote \mathbf{E}_i as the event that $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_i \in (\mathbb{Z}_q^n)^*$ and $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_i$ are linearly independent in \mathbb{Z}_q^n . We define \mathbf{E}_i^j (respectively \mathbf{D}_i^j) as the event that $\mathbf{u}_1, \mathbf{u}_2, \dots, \mathbf{u}_i \in (\mathbb{Z}_{q_j}^n)^*$ (respectively $(\mathbb{Z}_{p_j}^n)^*$) and these vectors are linearly independent in $\mathbb{Z}_{q_j}^n$ (respectively $\mathbb{Z}_{p_j}^n$) for $1 \leq i \leq n$ and $1 \leq j \leq k$. Our next goal is to compute $\Pr_{\mathbf{u}_i}[\mathbf{E}_i | \mathbf{E}_{i-1}]$ for all i where the probability is taken from $\mathbf{u}_i \leftarrow \mathbb{Z}_q^n$. We have the following claim and its proof is put to Appendix C.3.

Claim 5.3 We have

- If $1 \leq i \leq n - 1$, $\Pr_{\mathbf{u}_i}[\mathbf{E}_i | \mathbf{E}_{i-1}] \geq 1 - 2^{-n+i}$.
- $\Pr_{\mathbf{u}_n}[\mathbf{E}_n | \mathbf{E}_{n-1}] = \varphi(q)/q$, where φ is the Euler totient function.

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where $\mathbf{s} \stackrel{\S}{\leftarrow} \mathcal{S}$ and $\mathbf{e} \stackrel{\S}{\leftarrow} D_{\sigma\mathbf{B}}$.

Lemma A.29 (General Distributions [13]) *Let $\sigma > 0$ be a gaussian parameter, q be a modulo such that $q \geq \frac{\sigma}{\sqrt{\pi/\log(4n)}}$, \mathcal{S} be any distribution on \mathbb{Z}_q^n . Then it holds that*

$$\nu_\sigma(\mathcal{S}) \geq H_\infty(\mathcal{S}) - n \cdot \log(q/\sigma) - 1.$$

Lemma A.30 (Bounded Distributions [13]) *Let $\sigma > 0$ be a gaussian parameter, \mathcal{S} be a r -bounded (ℓ_2 norm) distribution on \mathbb{Z}_q^n . Then it holds that*

$$\nu_\sigma(\mathcal{S}) \geq H_\infty(\mathcal{S}) - \sqrt{2\pi n} \log(e) \cdot \frac{r}{\sigma}.$$

Then, we have the following reduction for the hardness of entropic RLWE [14].

Theorem A.31 *Assume that DSPR with parameter γ and RLWE with a B -bounded noise distribution χ holds. Let \mathcal{S} be a distribution such that $\nu_\sigma(\mathcal{S}) \geq n \log(\gamma \cdot \sqrt{n} \log(n)) + \omega(\log(\lambda))$ for some parameter σ . Then entSLWE for power-of-two cyclotomics with $\ell \geq 2n \log q + \omega(\log(\lambda))$ samples, secret distribution \mathcal{S} and error distribution Φ_{bin} is standard hard, where Φ_{bin} is defined as the distribution determined by choosing ℓ elements e_1, \dots, e_ℓ from Gaussian distribution χ with parameter $\sigma_0 \geq O(\sigma n \log(n) \sqrt{\ell} B)$ and $\mathbf{x} \stackrel{\S}{\leftarrow} \{0, 1\}^\ell$, and outputting $\sum_i x_i$.*

Remark A.32 *We note that the entropic RLWE considered in Corollary A.31 is with continuous error distribution over $K_{\mathbb{R}}/qR$. However, we require the error distribution of entropic RLWE to be discrete in our later application. Fortunately, there exists a simple reduction from entropic RLWE with continuous error distribution to the one with discrete error distribution by making use of the randomized rounding procedure in [42].*

Corollary A.33 *Assume that DSPR with parameter γ and RLWE with a B -bounded noise distribution χ holds. Let \mathcal{S} be a distribution such that $\nu_\sigma(\mathcal{S}) \geq n \log(\gamma \cdot \sqrt{n} \log(n)) + \omega(\log(\lambda))$ for some parameter σ . Then entSLWE for power-of-two cyclotomics with $\ell \geq 2n \log q + \omega(\log(\lambda))$ samples, secret distribution \mathcal{S} and error distribution Φ_{bin} defined as Theorem A.31 above for discrete Gaussian distribution χ with parameter $\sigma_0 \geq O(\sigma n \log^2(n) \sqrt{\ell} B)$ is standard hard.*

A.12 Leftover hash lemma

We will use the following two variants of the leftover hash lemma. Particularly, the first one is with respect to the case of \mathbb{Z} , and the second one is related to the case of \mathcal{O}_K .

Lemma A.34 (Particular case of Lemma 2.3 in [37]) *Let $m, n, q \in \mathbb{N}$ be integers and $\varepsilon \in (0, 1)$. Suppose \mathbf{s} is chosen from some distribution over \mathbb{Z}_q^n and $\mathbf{A} \stackrel{\S}{\leftarrow} \mathbb{Z}_q^{m \times n}$, $\mathbf{u} \stackrel{\S}{\leftarrow} \mathbb{Z}_q^m$ are chosen independently of \mathbf{s} from uniform distribution. Furthermore let Y be a random-variable (possibly) correlated with \mathbf{s} .*

This statement is no doubt true for $t = 0$. We assume that the statement holds for $t - 1$, from the noise growth property of homomorphic evaluation, for $i \in [5]$, we have

$$\begin{aligned}
 \|\mathbf{R}_{t,i}\|_\infty &\leq 2m(b-1)\|\mathbf{R}_{\text{var}(t)}\|_\infty + (1-x_{\text{var}(t)})\|\mathbf{R}_{t-1,\gamma_{t,i,0}}\|_\infty + x_{\text{var}(t)}\|\mathbf{R}_{t-1,\gamma_{t,i,1}}\|_\infty \\
 &\leq 2m^2\beta(b-1) + 2(t-1)m^2\beta(b-1) \tag{5} \\
 &= 2tm^2\beta(b-1)
 \end{aligned}$$

where (5) follows the induction assumption, the upper bound for $\|\mathbf{R}_{\text{var}(t)}\|_\infty$ and exactly one of $x_{\text{var}(t)}$ and $1 - x_{\text{var}(t)}$ is 1. Thus, the statement holds for t .

We choose the case $t = L \leq 4^d$ and complete the proof. \square

In [12, 27, 32], the authors utilized pseudorandom function PRF, where our lossy ABM-LTF tag (t_c, t_a) is defined by a PRF key $K \in \{0, 1\}^k$ such that $t_c = \text{PRF}_K(t_a)$. We encodes each bit of K , then the evaluation key ek consists of λ GSW encodings and both inversion key ik and tag key tk includes the PRF key K . In the evaluation (resp. inversion) step, we apply homomorphic evaluation in the public (resp. private) way from lemma B.7 to the PRF evaluation circuit.

However, there is no known construction for lattice-based PRF scheme in NC1 so that directly doing homomorphic computation to the PRF evaluation circuit results in a super-polynomial modulus q . In order to guarantee a polynomial modulus q , we need to apply the method in [30], which is using a fully homomorphic encryption scheme HE, evaluating the PRF circuit by HE.Eval instead of pubEval to get a HE ciphertext HE.ct related to the PRF value, and then evaluate HE.Dec on the GSW encodings of HE.sk and HE.ct by pubEval to get a GSW encoding of the PRF value. The homomorphic encryption scheme is defined as follows.

Definition B.9 (Fully Homomorphic Encryption [23]) HE is a special kind of PKE with an additional public evaluation key hevk generated in HE.Gen , an additional evaluation PPT algorithm HE.Eval and message space $\mathcal{M} = \{0, 1\}$:

- $\text{HE.Gen}(1^\lambda)$: Input a security parameter λ , output a public key hek , a public evaluation key hevk and a secret decryption key hdk .
- $\text{HE.Eval}(\text{hevk}, C, \text{ct}_1, \dots, \text{ct}_l)$: Input an evaluation key evk , homomorphically evaluate a circuit $f : \{0, 1\}^l \rightarrow \{0, 1\}$ on $\text{ct}_1, \dots, \text{ct}_l$, and output a ciphertext ct_f .

We require HE scheme has full homomorphism and compactness.

- **Full homomorphism.** For any boolean circuit C with polynomial depth $L = L(\lambda)$, and any messages $\text{msg}_1, \dots, \text{msg}_l \in \{0, 1\}$, we have

$$\Pr[\text{HE.Dec}(\text{hdk}, \text{HE.Eval}(\text{hevk}, C, \{\text{ct}_i\}_{i \in [l]})) = f(\text{msg}_1, \dots, \text{msg}_l)] = 1 - \text{negl}(\lambda),$$

where $(\text{hek}, \text{hevk}, \text{hdk}) \leftarrow \text{HE.Gen}(1^\lambda)$ and $\text{ct}_i \leftarrow \text{HE.Enc}(\text{hek}, \text{msg}_i)$.

- **Compactness.** The decryption circuit is independent of the evaluated circuit C .

choose of \mathbf{R}_{ent} , $\|\mathbf{R}\|_{\infty} \leq 2 \cdot 4^d \beta m^2 (b-1) + mb^2$. Therefore, for every column \mathbf{f} of \mathbf{F} , $\|\mathbf{R}\mathbf{f}\|_{\infty} \leq \|\mathbf{R}\|_{\infty} \|\mathbf{f}\|_{\infty} \leq m\beta(2 \cdot 4^d \beta m (b-1) + b^2) = \beta^*$ and we can model $\bar{\mathbf{F}}$ as a β^* -bounded distribution.

Finally, we apply the parameters $m^* = 2m$ and β^* to lemma 4.3 to get the lossiness $l = (\ell + \lambda) \log q + n \log p^*$. \square

Lemma B.15 (Indistinguishability) *Assume $\text{LWE}_{\ell,m,q,\chi}$ and $\text{HNFLWE}_{n,n,q,\chi}$ is hard, PRF is a secure pseudorandom function scheme and HE has IND-CPA security. The above ABM-LTF construction has indistinguishability.*

Proof. This proof is similar to the proof of indistinguishability of the ABM-LTF scheme from Libert et al. [32, Lemma 14].

\mathcal{A} is a PPT adversary to attack the indistinguishability property. We prove this theorem by hybrid arguments and the first game is the real indistinguishability game. Denote W_i as the event that adversary \mathcal{A} outputs 1 in hybrid i .

Hybrid 0: This hybrid is the experiment 0 in the indistinguishability game defined in B.6. The challenger generates the evaluation key ek as described in the ABM construction and gives ek to the adversary. The challenger answers each lossy tag queries by $\text{ABM.LTag}(\text{tk}, \cdot)$.

Hybrid 1: This hybrid is the same as hybrid 0 except that the challenger generates the evaluation key ek in a different way. Instead of applying the Lossy function to generate \mathbf{A} i.e. $\mathbf{A} \leftarrow \text{Lossy}(1^n, 1^m, 1^\ell, q)$, the challenger samples \mathbf{A} uniformly random on $\mathbb{Z}_q^{m \times n}$ i.e. $\mathbf{A} \stackrel{\$}{\leftarrow} \mathbb{Z}_q^{m \times n}$. From lemma 4.2, under the hardness of $\text{LWE}_{\ell,m,q,\chi}$, the lossy matrix and the uniform matrix are computationally indistinguishable, hence $|\Pr[W_0] - \Pr[W_1]| \leq n \text{Adv}_{\text{LWE}, \mathcal{A}_0}^{\ell,m,q,\chi}(\lambda)$ for some LWE indistinguishability adversary \mathcal{A}_0 .

Hybrid 2: This hybrid is the same as hybrid 1 except that the challenger samples $\{\mathbf{D}_i\}_{i \in [g]}$ uniformly at random i.e. $\mathbf{D}_i \stackrel{\$}{\leftarrow} \mathbb{Z}_q^{m \times m}$ for all $i \in [g]$, instead of making each \mathbf{D}_i be a GSW encoding of the bit hdk_i . From lemma 5.10, under the hardness of $\text{HNFLWE}_{n,n,q,\chi}$ and $m = kn \geq 2n + \omega(\log \lambda \cdot \log \log \lambda)$, we obtain that $|\Pr[W_1] - \Pr[W_2]| \leq mg \text{Adv}_{\text{pse}, \mathcal{A}_1}^{n,m,q}(\lambda)$ for some adversary \mathcal{A}_1 attacking the pseudorandomness in lemma 5.10.

Hybrid 3: This hybrid is the same as hybrid 2 except that the challenger encrypts k_0 bits of 0 by HE, i.e. $d_i \leftarrow \text{HE.Enc}(\text{hek}, 0)$ for $i \in [k_0]$, instead of encrypting each bit of the PRF key K . From the IND-CPA security of HE, $|\Pr[W_2] - \Pr[W_3]| \leq k_0 \text{Adv}_{\text{HE}, \mathcal{A}_2}^{\text{IND-CPA}}(\lambda)$ for some CPA security attacker \mathcal{A}_2 of HE.

Hybrid 4: This hybrid is the same as hybrid 3 except that the challenger answers the lossy tag queries by $\mathcal{O}_{\mathcal{T}_c}(\cdot)$, an oracle that returns a uniform random core tag $t_c \stackrel{\$}{\leftarrow} \mathcal{T}_c$, which substitute the oracle $\text{ABM.LTag}(\text{tk}, \cdot)$.

Next, we prove that for a secure PRF, the views of \mathcal{A} in Hybrid 3 and Hybrid 4 are computationally indistinguishable. We can use distinguishing ability in Hybrid 3 and Hybrid 4 of \mathcal{A} to construct a PRF attacker \mathcal{A}_3 . In detail, \mathcal{A}_3 interacts with a PRF challenger \mathcal{C} that uniformly choose a PRF key $K \stackrel{\$}{\leftarrow} \{0, 1\}^{k_0}$ and answer each query $M \in \{0, 1\}^{k_1}$ by returning either $\text{PRF}_K(M)$ or implementing a random function $R(\cdot)$ by lazy sampling and outputting $R(M)$. Initially,

as $\mathbf{C} \cdot \mathbf{s} \in (\mathbb{Z}_q^\ell)^*$ implies $\mathbf{s} \in (\mathbb{Z}_q^n)^*$, so it's remain to bound $\Pr_{\mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{s}} [|I| > \lambda \mid \mathbf{C} \cdot \mathbf{s} \in (\mathbb{Z}_q^\ell)^*]$. To do this, we denote the i -th column of the matrices \mathbf{B} and \mathbf{F} as \mathbf{b}_i and \mathbf{f}_i , respectively. We fix \mathbf{s} and \mathbf{C} such that $\mathbf{C} \cdot \mathbf{s} \in (\mathbb{Z}_q^\ell)^*$, and compute $\Pr[|I| > \lambda \mid \mathbf{C}, \mathbf{s}]$. For simplicity, we omit the condition \mathbf{C} and \mathbf{s} below as they are fixed now. Then, according to our definition of I , we have for every $i \in [m]$:

$$\begin{aligned} \Pr_{\mathbf{B}, \mathbf{F}} [i \in I] &= \Pr_{\mathbf{B}, \mathbf{F}} \left[[(\tilde{\mathbf{A}} \cdot \mathbf{s})_i]_p \neq [(\mathbf{B} \cdot \mathbf{C} \cdot \mathbf{s} + \frac{q}{p^*} \mathbf{F} \cdot \lfloor \mathbf{s} \rfloor_{p^*})_i]_p \right] \\ &= \Pr_{\mathbf{b}_i, \mathbf{f}_i} \left[\left\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \right\rangle + \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle + \frac{q}{p^*} \left\langle \mathbf{f}_i, \frac{p^*}{q} \mathbf{s} - \lfloor \mathbf{s} \rfloor_{p^*} \right\rangle \right] \neq \lfloor \langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle + \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle \rfloor_p \right] \\ &\leq \Pr_{\mathbf{b}_i, \mathbf{f}_i} \left[\left\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \right\rangle + \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle \in \text{border}_{p, q, \nu} \left(\left| \frac{q}{p^*} \left\langle \mathbf{f}_i, \frac{p^*}{q} \mathbf{s} - \lfloor \mathbf{s} \rfloor_{p^*} \right\rangle \right| \right) \right] \quad (8) \end{aligned}$$

$$\begin{aligned} &= \sum_{\tau} \Pr_{\mathbf{f}_i} \left[\left| \left\langle \mathbf{s} - \frac{q}{p^*} \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \right\rangle \right| = \tau \right] \Pr_{\mathbf{b}_i, \mathbf{f}_i} \left[\left\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \right\rangle + \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle \in \text{border}_{p, q, \nu}(\tau) \right] \\ &\leq \sum_{\tau} \Pr_{\mathbf{f}_i} \left[\left| \left\langle \mathbf{s} - \frac{q}{p^*} \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \right\rangle \right| = \tau \right] \cdot \frac{2\tau p}{q} \quad (9) \end{aligned}$$

$$\begin{aligned} &= \mathbf{E}_{\mathbf{f}_i} \left[\left| \left\langle \mathbf{s} - \frac{q}{p^*} \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \right\rangle \right| \right] \cdot \frac{2p}{q} \\ &\leq \frac{np\beta}{p^*} \leq \frac{1}{m}. \quad (10) \end{aligned}$$

where $\nu = \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle - \lfloor \frac{q}{p^*} \langle \lfloor \mathbf{s} \rfloor_{p^*}, \mathbf{f}_i \rangle \rfloor$ and (8) follows from the definition of a ‘border’, (9) follows by Lemma A.2 and the uniformity of $\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle$ over \mathbb{Z}_q , and (10) follows since each entry of $\mathbf{s} - \frac{q}{p^*} \lfloor \mathbf{s} \rfloor_{p^*}$ is bounded by $\frac{q}{2p^*}$ in absolute value, and each entry of \mathbf{f}_i is bounded by β in absolute value.

From the above, we have that $\mathbf{E}[|I|] = \sum_{i \in [m]} \mathbf{E}[i \in I] = \sum_{i \in [m]} \Pr[i \in I] \leq 1$. Furthermore for $i \in [m]$, the events $i \in I$ are mutually independent, as their probabilities are based on independently chosen \mathbf{b}_i 's and \mathbf{f}_i 's. Therefore, by the Chernoff bound, we have $\Pr_{\mathbf{B}, \mathbf{F}} [|I| > \lambda \mid \mathbf{C} \cdot \mathbf{s} \in (\mathbb{Z}_q^\ell)^*] < 2^{-\lambda}$, for any fixed \mathbf{s}, \mathbf{C} satisfying the condition. Using Equation (7) and the above calculation, we have

$$\Pr_{\mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{s}} [|I| > \lambda] < \delta + 2^{-\ell+1} + 2^{-\lambda}.$$

This completes the proof. \square

The bit-length of Z is $|I|(\log m + \log p)$, which is upper-bounded by $\lambda(\log m + \log p)$ with overwhelming probability, i.e., $1 - \varepsilon = 1 - (\delta + 2^{-\ell+1} + 2^{-\lambda})$. Therefore,

On the other hand, it's clear that $H_\infty(\mathbf{s}) \geq H_\infty(\mathbf{s} \mid [\mathbf{s}]_{q,p^*}) \geq (\ell + \lambda + 1) \cdot \log(q) + \omega(\log(\lambda))$. Then $\Pr[\mathbf{s} \notin (\mathbb{Z}_q^n)^*] = \Pr[\mathbf{s} = \mathbf{0}] \leq q^{-(\ell+\lambda+1)} \cdot 2^{-\omega(\log \lambda)} < q^{-\lambda}$, which means that δ can be set as $q^{-\lambda}$, and thus $\varepsilon = 2^{-\lambda} + q^{-\lambda} + 2^{-\ell+1}$.

Combining above with leftover hash lemma as Lemma A.34, we have

$$\left(\begin{array}{c} \tilde{\mathbf{A}} \\ \mathbf{a} \end{array} \right), \left[\begin{array}{c} [\tilde{\mathbf{A}}\mathbf{s}]_p \\ [\langle \mathbf{a}, \mathbf{s} \rangle]_p \end{array} \right] \overset{s}{\approx} \left(\begin{array}{c} \tilde{\mathbf{A}} \\ \mathbf{a} \end{array} \right), \left[\begin{array}{c} [\tilde{\mathbf{A}}\mathbf{s}]_p \\ [u]_p \end{array} \right].$$

The exact statistical distance is bounded by $2^{-\omega(\log \lambda)}$. Finally, we replace the lossy $\tilde{\mathbf{A}}$ with uniformly random \mathbf{A} to get:

$$\left(\begin{array}{c} \tilde{\mathbf{A}} \\ \mathbf{a} \end{array} \right), \left[\begin{array}{c} [\tilde{\mathbf{A}}\mathbf{s}]_p \\ [u]_p \end{array} \right] \overset{c}{\approx} \left(\begin{array}{c} \mathbf{A} \\ \mathbf{a} \end{array} \right), \left[\begin{array}{c} [\mathbf{A}\mathbf{s}]_p \\ [u]_p \end{array} \right].$$

Combining the above three hybrids proves (11).

By a simple hybrid argument as [1], we can further prove the desired statement

$$(\mathbf{A}, [\mathbf{A}\mathbf{s}]_p) \overset{c}{\approx} (\mathbf{A}, [u]_p).$$

□

C.3 Proof of Claim 5.3

To prove claim 5.3, we need the following estimation of sum of each prime's t -th power.

Lemma C.4 *Let $t \geq 2$ be a positive integer and $\{p_i\}_{i \geq 1}$ be all different sequenced primes. We have $\sum_{i=1}^{\infty} p_i^{-t} < 2^{-(t-1)}$.*

Proof. In the case $t \geq 3$,

$$\begin{aligned} \sum_{i=1}^{\infty} p_i^{-t} &< 2^{-t} + 3^{-t} + \sum_{j=2}^{\infty} (2j)^{-t} = 2^{-t} \cdot \left(1 + \left(\frac{3}{2}\right)^{-t} + \sum_{j=2}^{\infty} j^{-t} \right) \\ &= 2^{-t} \cdot \left(1 + \left(\frac{3}{2}\right)^{-t} + \int_1^{\infty} x^{-t} dx \right) \\ &= 2^{-t} \cdot \left(1 + \left(\frac{3}{2}\right)^{-t} + \frac{1}{t-1} \right) < 2^{-(t-1)}. \end{aligned}$$

For the case $t = 2$,

$$\begin{aligned} \sum_{i=1}^{\infty} p_i^{-2} &< \sum_{i=1}^{\infty} i^{-2} - 1 - \sum_{i=2}^{\infty} (2i)^{-2} \\ &= \sum_{i=1}^{\infty} i^{-2} - 1 - \frac{1}{4} \left(\sum_{i=1}^{\infty} i^{-2} - 1 \right) \\ &= \frac{3}{4} \left(\sum_{i=1}^{\infty} i^{-2} - 1 \right) = \frac{3}{4} \left(\frac{\pi^2}{6} - 1 \right) < 2^{-1}. \end{aligned}$$

□

Claim C.5 (Claim 5.3) *We have*

- If $1 \leq i \leq n-1$, $\Pr_{\mathbf{u}_i}[\mathbf{E}_i \mid \mathbf{E}_{i-1}] \geq 1 - 2^{-n+i}$.
- $\Pr_{\mathbf{u}_n}[\mathbf{E}_n \mid \mathbf{E}_{n-1}] = \varphi(q)/q$, where φ is the Euler totient function.

Proof. First, we can get a lower bound for each $\Pr_{\mathbf{u}_i}[\mathbf{E}_i^j \mid \mathbf{E}_{i-1}^j]$ where the probability is taken from $\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{q_j}^n$. For all $1 \leq j \leq k$,

$$\begin{aligned}
\Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{q_j}^n} [\mathbf{E}_i^j \mid \mathbf{E}_{i-1}^j] &= \Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{p_j}^n} [\mathbf{D}_i^j \mid \mathbf{D}_{i-1}^j] \\
&= \Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{p_j}^n} [\mathbf{u}_i \notin \text{span}\{\mathbf{u}_1, \dots, \mathbf{u}_{i-1}\} \mid \mathbf{D}_{i-1}^j] \\
&= 1 - \Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{p_j}^n} [\mathbf{u}_i \in \text{span}\{\mathbf{u}_1, \dots, \mathbf{u}_{i-1}\} \mid \mathbf{D}_{i-1}^j] \\
&= 1 - p_j^{-(n-i+1)}.
\end{aligned}$$

Since the k random variables $(\mathbf{u}_i \bmod q_j)$ for $j \in [k]$ is mutually independent when $\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_q^n$, we observe that for all $1 \leq i \leq n$,

$$\Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_q^n} [\mathbf{E}_i \mid \mathbf{E}_{i-1}] = \prod_{j=1}^k \Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_{q_j}^n} [\mathbf{E}_i^j \mid \mathbf{E}_{i-1}^j] = \prod_{j=1}^k (1 - p_j^{-(n-i+1)}).$$

If $1 \leq i \leq n-1$, by Lemma C.4 and union bound,

$$\Pr_{\mathbf{u}_i \xleftarrow{\S} \mathbb{Z}_q^n} [\mathbf{E}_i \mid \mathbf{E}_{i-1}] \geq 1 - \sum_{j=1}^k p_j^{-(n-i+1)} > 1 - 2^{-(n-i)}.$$

For the case $i = n$,

$$\Pr_{\mathbf{u}_n \xleftarrow{\S} \mathbb{Z}_q^n} [\mathbf{E}_n \mid \mathbf{E}_{n-1}] = \prod_{j=1}^k (1 - p_j^{-1}) = \frac{\varphi(q)}{q}.$$

□

C.4 Proof of Theorem 5.11 and Lemma 5.13

Proof (Lemma 5.13). It is easy to verify the correctness $\mathbf{T}_\mathbf{A} \cdot \mathbf{A} = \mathbf{G}$. The pseudorandomness of \mathbf{A} is directly based on the lemma 5.10. For the upper bound of the trapdoor's quality, since $\mathbf{R} \leftarrow \chi^{kn \times 2n}$, with overwhelming probability we have

$$\|\mathbf{T}_\mathbf{A}\|_\infty = \|\mathbf{R}\|_\infty + 1 \leq 2n\beta + 1.$$

□

Proof (Theorem 5.11). We prove the three algorithms TrapGen, LWEInvert and LWRInvert listed in Section 5.2 satisfy the properties.

From lemma 5.13, proof is done for TrapGen.

For the LWE samples $(\mathbf{A}, \mathbf{c} = \mathbf{A} \cdot \mathbf{s} + \mathbf{e})$, we compute

$$\mathbf{T}_{\mathbf{A}} \cdot \mathbf{c} = \mathbf{G} \cdot \mathbf{s} + \mathbf{T}_{\mathbf{A}} \cdot \mathbf{e}.$$

Note that $\|\mathbf{T}_{\mathbf{A}} \cdot \mathbf{e}\|_{\infty} \leq \|\mathbf{T}_{\mathbf{A}}\|_{\infty} \cdot \|\mathbf{e}\|_{\infty}$. With lemma 5.12, we can successfully output the secret \mathbf{s} from DecodeG($\mathbf{T}_{\mathbf{A}} \cdot \mathbf{c}$) if the error norm $\|\mathbf{e}\| \leq \frac{q}{2(b+1)\|\mathbf{T}_{\mathbf{A}}\|_{\infty}}$.

With LWR samples $(\mathbf{A}, \mathbf{c} = \lceil \mathbf{A} \cdot \mathbf{s} \rceil_p)$, the algorithm LWRInvert first transform c to the form of LWE samples

$$\mathbf{c}' = \left\lceil \frac{q}{p} \cdot \mathbf{c} \right\rceil = \left\lceil \frac{q}{p} \cdot \left\lceil \frac{p}{q} \cdot \mathbf{A} \cdot \mathbf{s} \right\rceil \right\rceil = \left\lceil \frac{q}{p} \cdot \left(\frac{p}{q} \cdot \mathbf{A} \cdot \mathbf{s} + \mathbf{e}' \right) \right\rceil = \mathbf{A} \cdot \mathbf{s} + \mathbf{e}$$

where $\mathbf{e}' \in (-1/2, 1/2]^m$ and $\mathbf{e} = \lceil (q/p)\mathbf{e}' \rceil = (q/p)\mathbf{e}' + \mathbf{e}''$ for some $\mathbf{e}'' \in (-1/2, 1/2]^m$. Hence $\|\mathbf{e}\|_{\infty} \leq (q/p + 1)/2 < q/p$. Therefore, as long as $p \geq 2(b+1)\|\mathbf{T}_{\mathbf{A}}\|_{\infty}$, the error norm $\|\mathbf{e}\|_{\infty} < \frac{q}{2(b+1)\|\mathbf{T}_{\mathbf{A}}\|_{\infty}}$ then DecodeG($\mathbf{T}_{\mathbf{A}} \cdot \mathbf{c}'$) recovers \mathbf{s} successfully. \square

C.5 Proof of Theorem 6.1

Theorem C.6 (ent-RLWE $_{\ell,q,\chi,S}$ to ent-sRLWR $_{q,p,\mathbf{B},\ell,S}$, Theorem 6.1) *Let $q \geq p \geq 2, n, \ell, B$ be positive integers such that $q \geq 18pB\ell n$, R be a ring of integers of a number field K with degree n , \mathbf{B} be a basis of R . Let χ be a B -bounded distribution over R with respect to basis \mathbf{B} , \mathcal{S} be a distribution over R_q^* . Then there exists a poly-time reduction from ent-RLWE $_{\ell,q,\chi,S}$ to ent-sRLWR $_{q,p,\mathbf{B},\ell,S}$*

Proof. Set $\beta = 2B$. Then the reduction can be obtained by the following two steps:

$$\text{ent-RLWE}_{\ell,q,\chi,S} \xrightarrow{(1)} \text{ent-RLWE}_{\ell,q,\chi+U_{\beta}(\mathbf{B}),S} \xrightarrow{(2)} \text{ent-sRLWR}_{q,p,\mathbf{B},\ell,S}.$$

The first reduction is straight-forward: given ℓ samples $(a_i, b_i) \in R_q \times R_q$ where $s \stackrel{\mathcal{S}}{\leftarrow}$, the reduction just adds independent samples from $U_{\beta}(\mathbf{B})$ to each b_i of the basis \mathbf{B} . It is easy to see the reduction maps the uniformly random distribution to itself, and $A_{s,\chi}$ to $A_{s,\chi+U_{\beta}}$, concluding the analysis of this part.

For the second reduction, we can bound the RD between samples from the two distributions. Thus, a solver of RLWR (as is) can be used to solve the Ring-LWE with the specified parameters.

Let \mathcal{X}_s be the distribution of a single ent-RLWR $_{q,p,\mathbf{B},S}$ sample, and let \mathcal{Y}_s be that of a single rounded RLWE $_{q,\chi+U_{\beta},S}$ sample under basis \mathbf{B} . By our setting of parameters $\beta = 2B$, the coefficients of $e \leftarrow \chi + U_{\beta}$ with respect to \mathbf{B} are B' -bounded, where $B' = 3B$. By the definition of Rényi divergence,

$$\begin{aligned} \text{RD}_2(\mathcal{X}_s || \mathcal{Y}_s) &= E_{a \leftarrow R_q} \frac{\Pr(\mathcal{X}_s = (a, \lfloor a \cdot s \rfloor_{\mathbf{B},p}))}{\Pr(\mathcal{Y}_s = (a, \lfloor a \cdot s \rfloor_{\mathbf{B},p}))} \\ &= E_{a \leftarrow R_q} \frac{1}{\Pr_{e \leftarrow \chi+U_{\beta}}(\lfloor a \cdot s + e \rfloor_{\mathbf{B},p} = \lfloor a \cdot s \rfloor_{\mathbf{B},p})}. \end{aligned}$$

Next we define the set

$$\text{border}_{q,p}(B') = \left\{ x \in \mathbb{Z}_q : x - \frac{q}{p} \left(\lfloor x \rfloor_p - \frac{1}{2} \right) < B' \text{ or } \frac{q}{p} \left(\lfloor x \rfloor_p + \frac{1}{2} \right) - x < B' \right\}.$$

For any $t \in \{0, \dots, n\}$, we define the set $\text{BAD}_{s,t} = \{a \in R_q : |\{i \in [n], (a \cdot s)_i \in \text{border}_{q,p}(B')\}| = t\}$, where $(a \cdot s)_i$ is the i th coefficient of $a \cdot s$ with respect to the basis \mathbf{B} . Fix t and $a \in \text{BAD}_{s,t}$, and below we do a case analysis:

- for any $i \in [n]$ such that $(a \cdot s)_i \notin \text{border}_{q,p}(B')$, we have

$$\Pr[\lfloor (a \cdot s)_i + e_i \rfloor_{\mathbf{B},p} = \lfloor (a \cdot s)_i \rfloor_{\mathbf{B},p}] = 1.$$

- For any $i \in [n]$ such that $(a \cdot s)_i \in \text{border}_{q,p}(B')$, the event $\lfloor (a \cdot s)_i + e_i \rfloor_{\mathbf{B},p} = \lfloor (a \cdot s)_i \rfloor_{\mathbf{B},p}$ holds at least in one of the two cases: (1) $e_i \in [-B', \dots, 0]$ or (2) $e_i \in [0, \dots, B']$.

Even though the coefficients of e (with respect to \mathbf{B}) might not be independent, and thus bounding $\Pr_{e \leftarrow \chi + U_\beta}(\lfloor a \cdot s + e \rfloor_{\mathbf{B},p} = \lfloor a \cdot s \rfloor_{\mathbf{B},p})$ is not straight-forward. To tackle this, we decompose $e = e' + e''$ where $e' \leftarrow \chi$ and $e'' \leftarrow U_\beta(\mathbf{B})$, and note that the coefficients of e'' dominates those of e' (as $\beta = 2B$). More importantly, the coefficients of e with respect to \mathbf{B} become independent of each others (in distribution) when we condition on e' . Since $a \in \text{BAD}_{s,t}$, a has exactly t coefficients in $\text{border}_{q,p}(B')$. Without loss of generality, we assume that the first t coefficients of $(a \cdot s) \in \text{border}_{q,p}(B')$, i.e., $(a \cdot s)_1, \dots, (a \cdot s)_t$. Next we would like to bound

$$\Pr[\lfloor a \cdot s + e \rfloor_{\mathbf{B},p} = \lfloor a \cdot s \rfloor_{\mathbf{B},p}] = \Pr[\forall i \in [t] \lfloor (a \cdot s)_i + e_i \rfloor_{\mathbf{B},p} = \lfloor (a \cdot s)_i \rfloor_{\mathbf{B},p}].$$

We know that for any $i \in [t]$, the event $\lfloor (a \cdot s)_i + e_i \rfloor_{\mathbf{B},p} = \lfloor (a \cdot s)_i \rfloor_{\mathbf{B},p}$ happens if $e''_i + e'_i$ falls belong to the correct half. Since e'_i is B bounded and $\beta = 2B$, this happens with probability at least $1/4$ over the choice of e''_i . As the $\{e''_i\}_{i \in [t]}$ are independent, we have

$$\Pr[\forall i \in [t] \lfloor (a \cdot s)_i + e_i \rfloor_{\mathbf{B},p} = \lfloor (a \cdot s)_i \rfloor_{\mathbf{B},p}] \geq (1/4)^t.$$

On the other hand, $s \in \mathcal{S} \subseteq (R_q)^*$, so $a \cdot s$ is uniformly random over R_q , implying $\Pr[a \in \text{BAD}_{s,t}] \leq \binom{n}{t} \cdot \left(1 - \frac{|\text{border}_{q,p}(B')|}{q}\right)^{n-t} \left(\frac{|\text{border}_{q,p}(B')|}{q}\right)^t$. Conditioning over the event $a \in \text{BAD}_{s,t}$, we have

$$\text{RD}_2(\mathcal{X}_s \| \mathcal{Y}_s) \leq \sum_{t=0}^n 4^t \cdot \Pr[a \in \text{BAD}_{s,t}] = \left(1 + \frac{3|\text{border}_{q,p}(B')|}{q}\right)^n.$$

By the definition of RD_2 , it is easy to see

$$\text{RD}_2(\mathcal{X}_s^\ell \| \mathcal{Y}_s^\ell) \leq \left(1 + \frac{3|\text{border}_{q,p}(B')|}{q}\right)^{\ell n},$$

It is clear that a' is uniformly random over R_q because a is uniformly random over R_q . On the other hand, b' can be written as

$$\begin{aligned} b' &= b + h + vg \\ &= \lfloor a \cdot s \rfloor_{\mathbf{B},p} + h + vg \\ &= \lfloor a' \cdot s - \frac{q}{p}v \cdot s \rfloor_{\mathbf{B},p} + h + vg \\ &= \lfloor a' \cdot s \rfloor_{\mathbf{B},p} + h + v(g - s). \end{aligned}$$

If $s \equiv g \pmod{\mathfrak{p}_i R}$, then by the Chinese Remainder Theorem A.10, $v(s - g) = 0 \pmod{pR}$. In this case, (a', b') is distributed according to $L_{s,q,p}^{i-1}(R, \mathbf{B})$. Otherwise if $s \not\equiv g \pmod{\mathfrak{p}_i R}$, we claim that $v(s - g) \pmod{\mathfrak{p}_i R}$ is uniformly random over $\mathfrak{p}_i R$ and is 0 mod all the other ideals $\mathfrak{p}_j R$'s for $j \neq i$: as R/\mathfrak{p}_i is a field, $v(s - g) \pmod{\mathfrak{p}_i R}$ is uniformly random for a random $v \pmod{\mathfrak{p}_i R}$, and any $(s - g) \not\equiv 0 \pmod{\mathfrak{p}_i R}$. Therefore, $v(g - s) + h$ is uniformly random mod $\mathfrak{p}_j R$ for all $j \leq i$, and is 0 mod all the remaining $\mathfrak{p}_j R$'s. Thus, the distribution of (a', b') follows $L_{s,q,p}^i(R, \mathbf{B})$ in this case, completing the proof. \square

C.8 Proof of Lemma 6.10

Lemma C.9 (Worst-case to average-case, Lemma 6.10) *Let $\mathcal{S}, \mathcal{S}_1, \mathcal{S}_2$ be distributions over R_q . For every $i \in \{1, \dots, g\}$, if $r \leftarrow \mathcal{S}_1$ is invertible with non-negligible probability, and $\text{RD}(\text{Coeff}_{\mathbf{B}}(\mathcal{S}_2) \| \text{Rot}_{\mathbf{B}}(s) \cdot \text{Coeff}_{\mathbf{B}}(\mathcal{S}_1)) \leq \text{poly}(\lambda)$ for any $s \in \text{Supp}(\mathcal{S})$. Then there exists a randomized poly-time reduction from worst-case (W) - D -RLWR $_{q,p,\mathbf{B},\ell,S}^i$ to average-case D -RLWR $_{q,p,\mathbf{B},\ell,S_2}^i$.*

Proof. Given sample $(a, b) \leftarrow L_{s,q,p}^i(R, \mathbf{B})$ for arbitrary $s \in \text{Supp}(\mathcal{S})$, the reduction transforms it into $(a', b') = (ar^{-1}, b + h) \in R_q \times R_p$, where $r \leftarrow \mathcal{S}_1$, and $h \in R_q$ is uniformly random mod $\mathfrak{p}_j R$ for all $j \leq \nu$ (where $\nu \leq i$), and 0 over mod all the other ideals. It's clear that a' is uniformly random, since a is uniformly random and r^{-1} is invertible. For the other term, we have $b' = b + h = \lfloor a \cdot s \rfloor_{\mathbf{B},p} + h = \lfloor ar^{-1} \cdot rs \rfloor_{\mathbf{B},p} + h$. Therefore, for all $s \in R_q$ and $i \in \{1, \dots, g\}$, this transformation maps $L_{s,q,p}^i(R, \mathbf{B})$ to $L_{s \cdot \mathcal{S}_1, q, p}^{\max\{\nu, i\}}(R, \mathbf{B})$.

Formally, the reduction is executed by repeating the following steps a polynomial number of times: Choose an r from \mathcal{S}_1 , and then estimate the acceptance probability of the oracle on the following two input distributions: the first is obtained from our input by applying the above transformation with parameters r , and $i - 1$; the second is obtained similarly using parameters r , and i . If in any of these polynomial number of attempts a non-negligible difference is observed between the two acceptance probabilities, output " $i - 1$ "; otherwise output " i ".

If the input distribution is $L_{s,q,p}^i(R, \mathbf{B})$, then in each of the attempts, the two distributions on which we estimate the oracle's acceptance probability are exactly the same, and we output " i " with overwhelming probability. If the input distribution is $L_{s,q,p}^{i-1}(R, \mathbf{B})$, we estimate the oracle's acceptance probability on $L_{s \cdot \mathcal{S}_1, q, p}^{i-1}(R, \mathbf{B})$ and $L_{s \cdot \mathcal{S}_1, q, p}^i(R, \mathbf{B})$.

according to our definition of I , we have for every $i \in [m]$:

$$\begin{aligned} \Pr_{\mathbf{B}, \mathbf{F}} [i \in I] &= \Pr_{\mathbf{B}, \mathbf{F}} \left[[(\tilde{\mathbf{A}} \cdot \mathbf{s})_i]_p \neq [(\mathbf{B} \cdot \mathbf{C} \cdot \mathbf{s} + \mathbf{F} \cdot (\mathbf{s} + [\mathbf{e}]))_i]_p \right] \\ &= \Pr_{\mathbf{b}_i, \mathbf{f}_i} \left[|[\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle + \langle \mathbf{s} + [\mathbf{e}], \mathbf{f}_i \rangle - \langle [\mathbf{e}], \mathbf{f}_i \rangle]_p \neq |[\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle + \langle \mathbf{s} + [\mathbf{e}], \mathbf{f}_i \rangle]_p \right] \\ &\leq \Pr_{\mathbf{b}_i, \mathbf{f}_i} [|\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle + \langle \mathbf{s} + [\mathbf{e}], \mathbf{f}_i \rangle \in \text{border}_{p, q, \nu}(|\langle \mathbf{e}, \mathbf{f}_i \rangle|)] \end{aligned} \quad (16)$$

$$\begin{aligned} &= \sum_{\tau} \Pr_{\mathbf{f}_i} [|\langle [\mathbf{e}], \mathbf{f}_i \rangle| = \tau] \cdot \Pr_{\mathbf{b}_i, \mathbf{f}_i} [|\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle + \langle \mathbf{s} + [\mathbf{e}], \mathbf{f}_i \rangle \in \text{border}_{p, q, \nu}(\tau)] \\ &\leq \sum_{\tau} \Pr_{\mathbf{f}_i} [|\langle [\mathbf{e}], \mathbf{f}_i \rangle| = \tau] \cdot \frac{2\tau p}{q} \end{aligned} \quad (17)$$

$$\begin{aligned} &= \mathbf{E}_{\mathbf{f}_i} [|\langle [\mathbf{e}], \mathbf{f}_i \rangle|] \cdot \frac{2p}{q} \\ &\leq \beta \cdot \|\mathbf{e}\|_1 \cdot \frac{2p}{q} \leq \beta \cdot \left(\|\mathbf{e}\|_1 + \frac{n}{2} \right) \cdot \frac{2p}{q} \end{aligned} \quad (18)$$

$$\leq \frac{np\beta(\sigma + 1)}{q} \leq \frac{1}{m}. \quad (19)$$

where $\nu = \langle \mathbf{s} + [\mathbf{e}], \mathbf{f}_i \rangle - \langle \mathbf{s} + \mathbf{e}, \mathbf{f}_i \rangle$ and (16) follows from the definition of a ‘‘border’’, (17) follows by Lemma A.2 and the uniformity of $\langle \mathbf{C} \cdot \mathbf{s}, \mathbf{b}_i \rangle$ over \mathbb{Z}_q , and (19) follows given the precondition $\|\mathbf{e}\| \leq n\sigma/2$, and each entry of \mathbf{f}_i is bounded by β in absolute value.

From the above, we have that $\mathbf{E}[|I|] = \sum_{i \in [m]} \mathbf{E}[i \in I] = \sum_{i \in [m]} \Pr[i \in I] \leq 1$. Furthermore for $i \in [m]$, the events $i \in I$ are mutually independent, as their probabilities are based on independently chosen \mathbf{b}_i ’s and \mathbf{f}_i ’s. Therefore, by the Chernoff bound, we have $\Pr_{\mathbf{B}, \mathbf{F}} [|I| > \lambda \mid \mathbf{C} \cdot \mathbf{s} \in (\mathbb{Z}_q^\ell)^*, \|\mathbf{e}\|_1 \leq n\sigma/2] < 2^{-\lambda}$, for any fixed $\mathbf{s}, \mathbf{C}, \mathbf{e}$ satisfying the condition. Using Equation (15) and the above calculation, we have

$$\Pr_{\mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{s}, \mathbf{e}} [|I| > \lambda] < \delta + 2^{-\ell+1} + 2^{-\lambda} + e^{-\left(\frac{\pi}{4} - \ln 2\right)n},$$

which completes the proof of Claim D.2. \square

The bit-length of Z is $|I|(\log m + \log p)$, which is upper-bounded by $\lambda(\log m + \log p)$ with overwhelming probability, i.e., $1 - \varepsilon = 1 - \left(\delta + 2^{-\ell+1} + 2^{-\lambda} + e^{-\left(\frac{\pi}{4} - \ln 2\right)n} \right)$. Therefore, we have

$$\begin{aligned} H_\infty^{\varepsilon' + \varepsilon} (f(\mathbf{s}) \mid \tilde{\mathbf{A}}, [\tilde{\mathbf{A}} \cdot \mathbf{s}]_p, \text{aux}) &\geq H_\infty^{\varepsilon' + \varepsilon} (f(\mathbf{s}) \mid \mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{C} \cdot \mathbf{s}, \mathbf{s} + [\mathbf{e}], Z, \text{aux}) \\ &\geq H_\infty^{\varepsilon'} (f(\mathbf{s}) \mid \mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{C} \cdot \mathbf{s}, \mathbf{s} + [\mathbf{e}], \text{aux}) - \lambda(\log m + \log p) \\ &\geq H_\infty^{\varepsilon'} (f(\mathbf{s}) \mid \mathbf{B}, \mathbf{C}, \mathbf{F}, \mathbf{s} + [\mathbf{e}], \text{aux}) - \ell \log q - \lambda(\log m + \log p) \\ &\geq H_\infty^{\varepsilon'} (f(\mathbf{s}) \mid \mathbf{s} + [\mathbf{e}], \text{aux}) - (\ell + \lambda) \log q \\ &\geq H_\infty^{\varepsilon'} (f(\mathbf{s}) \mid \mathbf{s} + \mathbf{e}, \text{aux}) - (\ell + \lambda) \log q. \end{aligned}$$

where the second and the third lines follow by Lemma A.7 and Claim D.2, and the last line follows by $q \geq mp$. This completes the proof of Lemma D.1. \square

according to D_σ , we can prove that $\Pr_{\mathbf{e}}[\sum_i |e_i| \geq nt] \leq 2^n \cdot \exp\left(-\frac{\pi nt^2}{\sigma^2}\right)$. Hence $\sum_i |e_i|$ has an upper bound $O(n\sigma)$ overwhelmingly, yielding that $\mathbf{E}[|\langle \mathbf{f}_i, \mathbf{e} \rangle|]$ can be bounded by $O\left(\frac{p\beta n\sigma}{q}\right)$. Therefore, q is with lower bound $O(nmp\beta\sigma)$. On the other hand, the lower bound itself of noise lossiness should also be considered. Specifically, for general entropic secrets, the noise lossiness is with lower bound $H_\infty(\mathbf{s}|\mathbf{s} + \mathbf{e})_g \geq H_\infty(\mathbf{s}) - n \log\left(\frac{q}{\sigma}\right) - 1$; for bounded entropic secrets, the noise lossiness is with lower bound $H_\infty(\mathbf{s}|\mathbf{s} + \mathbf{e})_b \geq H_\infty(\mathbf{s}) - \sqrt{2\pi n} \frac{r}{\sigma} \log e$, where r is the upper bound of ℓ_2 norm of secrets. For convenience of comparison, we relist the lower bound of q by our rounding lossiness approach, and the lower bounds of rounding lossiness for general and bounded entropic secrets as follows: $q > p^* \geq nmp\beta$, $H_\infty(\mathbf{s}|\lfloor \mathbf{s} \rfloor_{q,p^*})_g \geq H_\infty(\mathbf{s}) - n \log p^*$, $H_\infty(\mathbf{s}|\lfloor \mathbf{s} \rfloor_{q,p^*})_b \geq H_\infty(\mathbf{s}) - n \log\left(\frac{p^*}{q}\right)$, where γ is the upper bound of ℓ_∞ norm of secrets.

Based on these bounds, we can see that, for general entropic secrets, the modulus q of our framework is similar to the noise lossiness framework, if the two frameworks are required to have the same entropy lower bound. For bounded secrets, the modulus q of our framework can be saved at least a factor of $O(\sqrt{n})$ compared with the noise lossiness framework, if we minimize the entropy requirement of secrets of the two frameworks simultaneously. In other words, our approach can achieve better parameters than the noise lossiness approach. Besides, when working with certain secret distribution \mathcal{S} , rounding lossiness is easier to compute since the leakage is a determined function on \mathbf{s} .

To sum up, we find it more advantageous to work with our rounding lossiness framework compared with the *noise lossiness* presented in [13].

E Hardness of Entropic MLWR

Definition E.1 (Lossy Sampler over Ring, Definition 5.13 in [33]) Let n, ℓ, f, p, q, k be positive integers, $R = \mathcal{O}_K$ be the ring of integers of a field extension K with degree n , and ϕ be a distribution over $K_{\mathbb{R}}$. We define the following efficient lossy sampler $\tilde{\mathbf{A}} \xleftarrow{\$} \text{Lossy}(1^n, 1^\ell, 1^f, 1^k, q, \phi)$ as:

$\text{Lossy}(1^n, 1^\ell, 1^f, 1^k, q, \phi) : \text{Sample } \mathbf{D} \xleftarrow{\$} (R_q)^{\ell \times f}, \mathbf{C} \xleftarrow{\$} (R_q)^{f \times k}, \mathbf{F} \xleftarrow{\$} \phi^{\ell \times k}$
and output $\tilde{\mathbf{A}} = \mathbf{D} \cdot \mathbf{C} + \mathbf{F}$.

The output of Lossy algorithm is computationally indistinguishable from uniformly random sample according to the following lemma and corollary.

Lemma E.2 Let $\mathbf{A} \xleftarrow{\$} (R_q)^{\ell \times k}$, and let $\tilde{\mathbf{A}} \xleftarrow{\$} \text{Lossy}(1^n, 1^\ell, 1^f, 1^k, q, \phi)$. Then, according to the module-RLWE $_{\ell, f, q, \phi}$ assumption, we have: $\mathbf{A} \stackrel{c}{\approx} \tilde{\mathbf{A}}$.

Corollary E.3 Adopt the notations in Lemma A.24 and Lemma E.2. Assuming the RLWE $_{\ell, q, \phi}$ problem is computationally hard, then $\mathbf{A} \stackrel{c}{\approx} \tilde{\mathbf{A}}$.

We denote that vector $\mathbf{r} \in (R)^k$ maximal belongs to a factor \mathcal{I} of qR , abbreviated as $\mathbf{r} \in_{\max} \mathcal{I}R$ if the following conditions hold.

Proof. We divide the event $|I| > \lambda$ into two cases: $\mathbf{s} \in_{\max} \langle 1 \rangle$ and $\mathbf{s} \notin_{\max} \langle 1 \rangle$:

$$\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda] = \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \notin_{\max} \langle 1 \rangle] \cdot \Pr_{\mathbf{s}}[\mathbf{s} \notin_{\max} \langle 1 \rangle] \quad (22)$$

$$+ \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle] \cdot \Pr_{\mathbf{s}}[\mathbf{s} \in_{\max} \langle 1 \rangle] \quad (23)$$

$$\leq \Pr_{\mathbf{s}}[\mathbf{s} \notin_{\max} \langle 1 \rangle] + \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle] \quad (24)$$

$$< \delta + \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle]. \quad (25)$$

It remains to bound $\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle]$. In order to compute $\Pr[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle]$, we divide the condition into two cases: $\mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle$ and $\mathbf{C} \cdot \mathbf{s} \notin_{\max} \langle 1 \rangle$. Then we have:

$$\begin{aligned} & \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle] \\ &= \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle] \cdot \Pr_{\mathbf{C}, \mathbf{s}}[\mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle \mid \mathbf{s} \in_{\max} \langle 1 \rangle] \\ &+ \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \notin_{\max} \langle 1 \rangle] \cdot \Pr_{\mathbf{C}, \mathbf{s}}[\mathbf{C} \cdot \mathbf{s} \notin_{\max} \langle 1 \rangle \mid \mathbf{s} \in_{\max} \langle 1 \rangle]. \end{aligned}$$

Combining equations 22, we have

$$\begin{aligned} \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda] &< \delta + \Pr_{\mathbf{C}, \mathbf{s}}[\mathbf{C} \cdot \mathbf{s} \notin_{\max} \langle 1 \rangle] + \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle] \\ &< \delta + \frac{n}{2^{f-1}} + \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle]. \end{aligned} \quad (26)$$

Finally, it remains to bound $\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle]$. It is easy to verify that $\mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle$ implies $\mathbf{s} \in_{\max} \langle 1 \rangle$, so $\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{s} \in_{\max} \langle 1 \rangle, \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle] = \Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle]$. On the other hand, for uniformly random matrix $\mathbf{D} \in (R_q)^{\ell \times f}$ and $\mathbf{C}\mathbf{s} \in_{\max} \langle 1 \rangle$, it's easy to verify $\mathbf{D} \cdot \mathbf{C}\mathbf{s}$ is uniformly at random by similar arguments as Theorem 5.7 in [33].

Now we claim that $\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle] < 2^{-\lambda}$. To show this, we fix any \mathbf{s} and \mathbf{C} such that $\mathbf{C} \cdot \mathbf{s} \in_{\max} \langle 1 \rangle$, and compute $\Pr_{\mathbf{D}, \mathbf{C}, \mathbf{F}, \mathbf{s}}[|I| > \lambda \mid \mathbf{C}, \mathbf{s}]$. For simplicity, below we omit the condition \mathbf{s}, \mathbf{C} as they are fixed now.

We denote the i -th row of the matrices \mathbf{D} and \mathbf{F} as \mathbf{d}_i and \mathbf{f}_i , respectively, and omit the common basis \mathbf{B} in the subscript of rounding function for simplicity.

The exact statistical distance is bounded by $2^{-\omega(\log \lambda)}$. Finally, we replace the lossy $\tilde{\mathbf{A}}$ with uniformly random \mathbf{A} to get:

$$\left(\begin{array}{c} [\tilde{\mathbf{A}}] \\ [\mathbf{a}] \end{array}, \begin{array}{c} [[\tilde{\mathbf{A}}\mathbf{s}]_{\mathbf{B},p}] \\ [u]_{\mathbf{B},p} \end{array} \right) \stackrel{c}{\approx} \left(\begin{array}{c} [\mathbf{A}] \\ [\mathbf{a}] \end{array}, \begin{array}{c} [[\mathbf{A}\mathbf{s}]_{\mathbf{B},p}] \\ [u]_{\mathbf{B},p} \end{array} \right).$$

Combining the above three hybrids proves (31). By a simple hybrid argument as [1], we can further prove the desired statement

$$(\mathbf{A}, [\mathbf{A}\mathbf{s}]_{\mathbf{B},p}) \stackrel{c}{\approx} (\mathbf{A}, [u]_{\mathbf{B},p}).$$

□