Vectorial Decoding Algorithm for Fast Correlation Attack and Its Applications to Stream Cipher Grain-128a

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Abstract. Fast correlation attacks, pioneered by Meier and Staffelbach, is an important cryptanalysis tool for LFSR-based stream cipher, which exploits the correlation between the LFSR state and key stream and targets at recovering the initial state of LFSR via a decoding algorithm. In this paper, we develop a vectorial decoding algorithm for fast correlation attack, which is a natural generalization of original binary approach. Our approach benefits from the contributions of all correlations in a subspace. We propose two novel criterions to improve the iterative decoding algorithm. We also give some cryptographic properties of the new FCA which allows us to estimate the efficiency and complexity bounds. Furthermore, we apply this technique to well-analyzed stream cipher Grain-128a. Based on a hypothesis, an interesting result for its security bound is deduced from the perspective of iterative decoding. Our analysis reveals the potential vulnerability for LFSRs over generic linear group and also for nonlinear functions with biased multidimensional linear approximations such as Grain-128a.

Keywords: Linear Approximation · Fast Correlation Attack · Iterative Decoding · Grain-128a.

1 Introduction

Stream ciphers are a widely used class of symmetric-key cryptosystem. A key stream sequence is generated from the initial state derived from the key. The plaintext is encrypted by XORing with the key stream in the same length.

Linear feedback shift register (LFSR) based stream ciphers form an important class of stream cipher system, in which one or more LFSRs are often used. LFSRs could be defined over different algebraic structures, such as finite fields and generic linear group. Besides for LFSR, these ciphers usually adopt a nonlinear filter function or a finite state automata (FSM) with nonlinear update function. The history of these ciphers can be traced back to decades ago, e.g., LILI-128 [CDF+02], the SNOW family [EJ00, EJ03, UEA06, EJMY19] and the Grain family etc.

The Grain family includes three well-known stream ciphers: Grain-128a [ÅHJM11], Grain-128 [HJMM06] and Grain-v1 [HJM07]. Grain-v1 is in the eSTREAM portfolio and Grain-128a is standardized by ISO/IEC [29115]. All the members of the Grain family share a similar structure. Several lightweight ciphers proposed recently also adopt similar

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structures [AM15, AHMN13, MAM16]. An important attack for Grain-v1 is near collision attack [ZLFL13], which is improved in [ZXF18]. Since Grain-128 adopts quadratic function, the dynamic cube attack plays an important role in its cryptanalysis [DS11]. To avoid the dynamic cube attack, Grain-128a adopts a nonlinear function with higher degree. However, the Grain family is reported to be vulnerable for fast correlation attacks (FCA) in CRYPTO 18 [TIM+18].

FCA is pioneered by Meier and Staffelbach in 1989 [MS89]. Generally speaking, FCA exploits the correlation between the key stream and the state or the outputs of LFSR. The problem of recovering initial state of LFSR is transformed into a decoding problem. The linear part of the stream cipher is treated as a linear code, and the nonlinear part of the stream cipher is treated as noise. According to the differences of decoding strategies, these FCA approaches can be roughly divided into two classes.

The first class adopts one-pass decoding algorithm. For example, the FCA adopts convolution codes and Viterbi decoding algorithm [JJ99b], which is improved it by turbo codes [JJ99a]. Another FCA adopts maximum likelihood decoding on a reduced set of information bits [CJS00]. The parity-checks are usually folded to eliminate partial bits. List decoding and polynomial reconstruction can also be applied in FCA [MFI02, JJ00]. An important improvement is accelerating the parity-check evaluations by fast Walsh-Hadamard transform (FWHT) [CJM02]. This technique is applied in cryptanalysis of the stream cipher E0 [LV04]. It was later generalized to extension fields and applied to stream cipher SNOW 2.0 [ZXM15]. A recent improvement of FCA is based on commutative property and applied to Grain family [TIM+18].

The second class adopts several-pass decoding algorithm. After Meier and Staffelbach’s original FCA, low-density parity-check code (LDPC) is introduced into FCA to improve the iterative decoding algorithm [CT00]. There are many related works in this area, such as [ÅLHJ12, CT00, Gol01, CGD96, GH05, MG91, MG93]. Intuitively, iterative decoding algorithm seems to be more powerful, as their decoding abilities are closer to Shannon’s bound. However, comparing with the FCA decoding by information set, it is usually very hard to describe its cryptographic properties by mathematical language, and also lacks of a convenient approach to work on extension fields. Thereby, its direct application to modern stream ciphers is very limited.

Our Contributions.

In this paper, we propose a vectorial iterative decoding algorithm for fast correlation attack, which generalizes Meier and Staffelbach’s original FCA very naturally. Our approach benefits from the contributions of all correlations in a subspace and thereby more powerful than the binary version. We propose two novel criterions to improve the iterative decoding algorithm and perform a scaled experiments to verify its validity. We also give some cryptographic properties for the first iteration, which allows us to estimate the efficiency and complexity bound via probability distribution approximations.

Furthermore, we apply it to the well-analyzed stream cipher Grain-128a. Based on a hypothesis that the initial probability distribution of noises is close to symmetric probability distribution, and there exist parity-checks with two taps or with special form, we give a data complexity bound estimation in the sense of being able to correct errors of the noisy sequence. The result shows maybe its potential security bound is lower than we thought from the perspective of vectorial iterative decoding. Our analysis reveals the potential vulnerability for LFSRs over generic linear group and also for nonlinear functions with biased multidimensional linear approximations such as Grain-128a.
Outline.

The rest of the paper is organized as follows. Section 2 is preliminary. The details of vectorial decoding algorithm are described in section 3. In section 4, we propose some cryptographic properties and perform an scaled experiment. How to apply the new FCA to Grain-128a is explained in section 5. Section 6 consists of some further problems. Finally, we conclude the paper.

2 Preliminary

2.1 Notations and Definitions

Some notations are introduced for convenience.

- Given 2 binary row vectors \( \mathbf{x} = (x_1, \cdots, x_{n-1}) \in \mathbb{F}_2^n \) and \( \mathbf{y} = (y_1, \cdots, y_{n-1}) \in \mathbb{F}_2^n \), their inner product is denoted by \( \mathbf{x} \cdot \mathbf{y} = \oplus_{i=0}^{n-1} x_i y_i \). The Hamming weight of \( \mathbf{x} \) are denoted by \( w(\mathbf{x}) \).

- Let \( F : \mathbb{F}_2^n \rightarrow \mathbb{F}_2^n \) denote a vectorial Boolean function. A binary linear approximation of \( F \) with \( m \)-bit input mask \( \mathbf{u} = (u_1, \cdots, u_m) \) and \( n \)-bit output mask pair \( \mathbf{v} = (v_1, \cdots, v_n) \) can be represented by \( \mathbf{u} \cdot \mathbf{x} \oplus \mathbf{v} \cdot F(\mathbf{x}) \). When we have \( 1 \leq r \leq m+n \) linearly independent mask pair \( (u_1, v_1), \cdots, (u_r, v_r) \), a vectorial (or multidimensional) linear approximation is denoted by \( U \mathbf{x} \oplus V F(\mathbf{x}) \), where the \( i \)-th row of \((U,V)\) is \((u_i,v_i)\). \( \mathbf{x} \) are treated as a column vector unless otherwise stated.

- Linear correlation is used to measure the bias of a binary linear approximation. Let \( c(\mathbf{x}) = \mathbf{u} \cdot \mathbf{x} \oplus \mathbf{v} \cdot F(\mathbf{x}) \), the correlation of the binary linear approximation is defined by \( c(\mathbf{u}, \mathbf{v}) = c(e) = 2^{-m} (\# \{ \mathbf{x} : e(\mathbf{x}) = 0 \} - \# \{ \mathbf{x} : e(\mathbf{x}) = 1 \} ) \). Similarly, let \( e(\mathbf{x}) = U \mathbf{x} \oplus V F(\mathbf{x}) \), \( \mathbf{w} \) is an \( r \) bits binary linear mask, the correlation of linear approximation with mask pair \((\mathbf{w}U, \mathbf{w}V)\) is \( c(\mathbf{w}) = 2^{-m}(\# \{ \mathbf{x} : \mathbf{w} \cdot e(\mathbf{x}) = 0 \} - \# \{ \mathbf{x} : \mathbf{w} \cdot e(\mathbf{x}) = 1 \} ) \).

- Let \( X \sim P \) denote a discrete random variable follows distribution \( P \) and takes values in \( \mathbb{F}_2^n \). Its probability density function \( p(\mathbf{x}) \) is denoted by \( (p_{(0,\cdots,0)}, \cdots, p_{(1,\cdots,1)}) = \Pr(\mathbf{X} = (0, \cdots, 0)), \cdots, \Pr(\mathbf{X} = (1, \cdots, 1)) \).

- Let \( a \in \mathbb{F}_2^n \) denote a binary vector. There is an integer \( a = \sum_{i=0}^{m-1} a_i 2^i \) corresponding to \( a \). For convenience, we alternatively use them if there is no ambiguity in the context, especially as a subscript. For example, for a probability density function \( (p_{(0,\cdots,0)}, \cdots, p_{(1,\cdots,1)}) \), we mean the same thing when denote it by \((p_0, \cdots, p_{2^m-1})\).

- Let \( M_m(\mathbb{F}_2) \) denote the \( m \times m \) matrix ring over \( \mathbb{F}_2 \). Given a LFSR with rank \( d \) and \( m \)-bit cell, its generator is denoted by \( L(\mathbf{x}) = E + C_1 x + C_2 x^2 + \cdots + C_d x^d \in M_m(\mathbb{F}_2)[\mathbf{x}] \), where \( C_d \) is nonsingular and \( E \) is the identity matrix. The number of information bits of \( L(\mathbf{x}) \) are denoted by \( k = d \times m \). If \( L(\mathbf{x}) \in \mathbb{F}_2^{2m} [\mathbf{x}] \), it can also be mapped into \( GL_m(\mathbb{F}_2)[\mathbf{x}] \).

- Give 2 positive integers \( a \) and \( b \) with \( gcd(a,b) = 1 \). The \( b \)-cyclotomic coset modulo \( a \) containing \( i \) is denoted by \( C_i = \{ i, ib, \cdots, ib^{r-1} \} \mod a \), where \( r \) is the smallest positive integer such that \( ib^r \equiv i \mod a \). The minimal integer in \( C_i \) is called coset header and denoted by \( \bar{i} \). All coset headers form a set \( \mathcal{R}_{b,a} \).

- Given 2 vectors \( \mathbf{a} = (a_1, \cdots, a_n) \in \mathbb{R}^n \) and \( \mathbf{b} = (b_1, \cdots, b_n) \in \mathbb{R}^n \). The notation \( \mathbf{a} \succ \mathbf{b} \) implies that there is at least one \( 1 \leq j \leq n \) satisfying \( a_j > b_j \), while \( \preceq \) has reverse meaning.
Walsh-Hadamard Transform

Walsh-Hadamard transform is a spectral tool widely used in cryptanalysis of linear type. Let $X \sim P$ denote a discrete random variable which take values in $\mathbb{F}_2^n$. The Walsh-Hadamard transform of $X$ is defined by

$$W(X)_w = 2^{-m} \sum_{x \in \mathbb{F}_2^n} p_x (-1)^{w \cdot x}.$$ 

Since Walsh-Hadamard transform is a linear operator for XOR, let random variable $X = X_1 \oplus X_2 \oplus \cdots \oplus X_k$, we can efficiently compute probability distribution of $X$ with the help of the convolution property

$$p_x = W^{-1}(W(X_1) \times \cdots \times W(X_k))_x.$$ 

Square Euclid Imbalance

Relative entropy (or Kullback–Leibler divergence) is used to measure the difference between two probability distributions $P$ and $Q$, i.e.,

$$D(p(x) \parallel q(x)) = \sum_x p_x \log \frac{p_x}{q_x}.$$

If $p(x)$ is close to $q(x)$, i.e., $p_x = q_x + \epsilon(x)$, the relative entropy could be approximated by $D(p(x) \parallel q(x)) \approx \frac{1}{2} \sum_x \frac{(p_x - q_x)^2}{q_x} + O(\epsilon^3(x))$. The summation term is usually called capacity, and denoted by $C(p \parallel q)$. Square Euclid imbalance (SEI) is defined to be the capacity between a probability distribution and uniform distribution, i.e.,

$$\Delta(p(x)) = 2^m \sum_x (p_x - \frac{1}{2^m})^2$$

The following theorem reveals the relationship between SEI and linear correlation.

Theorem 1 ([BJV04]). Let $X \in \mathbb{F}_2^n$ be a random variable with density function $p_x$, then its SEI

$$\Delta(p(x)) = \sum_w \epsilon^2(w) = \sum_{w \neq 0} \epsilon^2(w),$$

where $\epsilon(x) = p_x - 2^{-m}$, $\epsilon(w)$ denotes the FWHT of $\epsilon(x)$. For convenience, we use $\Delta(p)$ if $x$ is well known in the context, or $\Delta(X)$ if the random variable $X$ with density function $p(x)$ is clear. Particularly, we have $\epsilon^2(e) = \Delta(p)$ when $m = 1$. 

Parity-Check and Characteristic Polynomial

A parity-check corresponds to an equation which fulfills the LFSR output sequence $x_t$. For example, it is well known that any multiples of $L(x) \in \mathbb{F}_{2^m}[x]$ is a parity-check. Usually, only those very sparse parity-checks with low degree are exploited in FCA.

Let set $H(\tau + 1, d)$ denote all parity-checks with $\tau + 1$ taps and degree $d$, abbreviated by $H$ without ambiguity. The available parity-checks at position $n$ denoted by $H(n) \subseteq H$. Suppose a parity-check for sequence $x_t$ is denoted by

$$G_n x_t + \cdots + G_1 x_{t+n-1} + E x_{t+n} = 0,$$
where $G_n$ is nonsingular. Its characteristic polynomial is denoted by $F_n(x) = \det(Ex + A) = \det(\sum_{i=0}^{n} G_{n-i}x^i)$, where $A$ denotes the companion matrix

$$A = \begin{pmatrix}
0 & E & 0 & 0 & \cdots & 0 \\
0 & 0 & E & 0 & \cdots & 0 \\
\vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\
0 & 0 & 0 & \cdots & 0 & E \\
G_n & G_{n-1} & G_{n-2} & \cdots & G_2 & G_1
\end{pmatrix}.$$ 

### 2.2 A Brief Description of Original FCA

Meier and Staffelbach’s original FCA includes a precomputation phase and a decoding phase.

**Precomputation Phase**

Let LFSR’s generator polynomial $L(x) \in F_2[x]$. The purpose of precomputation phase is finding sufficient very sparse parity-checks with low degree, which is a hard open problem. One way recommended by Zeng [ZYR91] is evaluating logarithms in finite fields of characteristic 2. It is rather efficient to find low weight multiples, but the degree is not promised to be low. Another way is by extended K-tree algorithm based on general birthday collision [NS15]. The extended k-tree algorithm can be used to find low weight multiples of polynomial with not so large degree with flexible parameters.

**Decoding Phase**

The decoding phase targets to recover the initial state of LFSR from key stream. Suppose we have found sufficient suitable parity-checks $x_n \oplus a_n^{(i)} = 0$, where $a_n^{(i)}$ is the sum of $\tau$ taps $a_n^{(i)} = \sum_{k=1}^{\tau} x_{n-ik}$. The check value is $z_n \oplus b_n^{(i)}$, where $b_n^{(i)} = \sum_{k=1}^{\tau} z_{n-ik}$ is the sum of $\tau$ key stream bits corresponding to $x_{n-ik}$. The nonlinear part of a stream cipher is modeled as a binary symmetric channel (BSC), the crossover probability is $p = Pr[x_n \oplus z_n = 1]$. The critical part of decoding phase is calculating a posteriori probability (APP) with priori distribution symbol by symbol. Suppose that the check values are all 0 for a subset $H_0 \subseteq H$, then by Bayes’ formula,

$$p^* = \frac{p \prod_{i \in H_0} (1 - s_i) \prod_{i \in H \setminus H_0} s_i}{p \prod_{i \in H_0} (1 - s_i) \prod_{i \in H \setminus H_0} s_i + (1 - p) \prod_{i \in H \setminus H_0} (1 - s_i) \prod_{i \in H_0} s_i}$$

where each $s_i = s(p_{i_1}, \ldots, p_{i_\tau}) = Pr[a_n^{(i)} = b_n^{(i)}]$ depends on the probability of $\tau$ symbols involved in parity-check. Moreover, $s_i$ can be calculated recursively in the BSC Model

$$s(p_{i_1}, \ldots, p_{i_\tau}) = s(p_{i_1}, \ldots, p_{i_{\tau-1}}) + (1 - s(p_{i_1}, \ldots, p_{i_{\tau-1}}))(1 - s(p_{i_1}, \ldots, p_{i_{\tau-1}}))$$

The specific process is depicted in Algorithm 1. For more details we refer to the original paper [MS89].

### 3 Fast Correlation Attack Based on Vectorial Iterative Decoding Algorithm

#### 3.1 Channel Model

Our channel model is symmetric channel (SC) instead of discrete memoryless channel (DMC). The received word is the transmitted word XOR noise, i.e., $z = x \oplus e$. A symmetric
Algorithm 1 Meier and Staffelbach’s binary iterative decoding Algorithm B

**Input:** A key stream sequence \( z \) of length \( N \) and \( H \).

1. Calculate the probability threshold \( p_{\text{thr}} \) and quantity threshold \( N_{\text{thr}} \).
2. For round \( r \in \{1, 2, \cdots \} \)
3. For iteration \( i \) from 1 to a small integer
4. Calculate \( \text{APP}_p^* \) from priori probability \( p \), assign \( p_n^* = p_n \) for all position \( n \).
5. If \( N_w \geq N_{\text{thr}} \) where \( N_w = |\{n| p_n > p_{\text{thr}}\}| \), break;
6. Complement the bits of \( z \) with \( p_n > p_{\text{thr}} \).
7. Reset all positions to initial probability \( p \).
8. If \( z \) satisfies all parity-checks, break.
9. Terminate with \( x = z \).

Figure 1: Channel Model for VFCA

The channel model has a transition matrix

\[
M = \begin{pmatrix}
\Pr(z_1|x_1) & \Pr(z_2|x_1) & \cdots & \Pr(z_m|x_1) \\
\Pr(z_1|x_2) & \Pr(z_2|x_2) & \cdots & \Pr(z_m|x_2) \\
\vdots & \vdots & \ddots & \vdots \\
\Pr(z_1|x_m) & \Pr(z_2|x_m) & \cdots & \Pr(z_m|x_m)
\end{pmatrix}
\]

Each row is a permutation of another row, and so as to columns. Moreover, the sum of each row equals 1 as the definition of SC. Symmetric channel can be treated as an extended BSC. Its channel capacity is certainly \( C = m - H(r) \), where \( r \) denotes a row of \( M \).

Suppose we have a linear approximation with dimension \( m \), i.e.,

\[
\bigoplus_{i \in \{1, \cdots, #T_x\}} U_i x_{j(i)} \oplus \bigoplus_{j \in \{1, \cdots, #T_z\}} V_i z_{j(i)} = e.
\]  

Similarly as BSC, the channel noise vector \( e \) is XORed to \( \bigoplus_{i \in \{1, \cdots, #T_x\}, j(i) \in T_x} U_i x_{j(i)} \), and the output is \( \bigoplus_{j \in \{1, \cdots, #T_z\}, i \in T_z} V_i z_{j(i)} \), see Fig. 3.1.

**Remark 1.** When we are discussing a generic multidimensional linear approximation, we can always obtain a linear approximation with form \( U'x' \oplus V'z' \), i.e., only including one input vector \( x' \) and one output vector \( z' \), for example, by rewriting \( x' \) to a larger input vector of dimension \( m \times \#T_x \). Thus the rank of \( U \) becomes larger than those \( U_i \). However, despite that we are interesting to those linear approximations with large dimension and large SEI, the SEI is hard to always increase sufficiently as the dimension increases. Thus we pick multidimensional linear approximation with form (3) as an generic form.

### 3.2 Checking Parity with Vectorial Noise

Let \( l \in H^{(n)} \) denote a specific check equation:

\[
l : E x_n \oplus G_1 x_{n-1} \oplus \cdots \oplus G_n x_{n-d} = 0.
\]
In order to parity-check over matrix ring, these $G_1, \ldots, G_n$ are restricted by those matrices $U_i$ in (3). More specifically, we require that all $U_i$ are nonsingular. For each $G_k$, all $U_i$ satisfy $U_i(G_k U_i^{-1} = G_k'$, which implies that if $U_i, U_j$ satisfies $U_i G_k U_i^{-1} = G_k'$ and $U_j G_k U_j^{-1} = G_k'$, then $(U_j^{-1}U_i)G_k(U_j^{-1}U_i)^{-1} = G_k$, i.e., $U_j^{-1}U_i \in C(G_k)$, where $C(G_k)$ denotes the centralizer of $G_k$ in $GL_m(\mathbb{F}_2)$.

For a parity-check $l$ we could multiply it with $U_1, U_2, \ldots, U_{\#T_e}$ respectively,

$$U_i(E \delta_{n+j(i)} \oplus G_1 \delta_{n+1+j(i)} \oplus \cdots \oplus G_d \delta_{n-d+j(i)}) = 0.$$ 

Thus we have

$$E(U_i \delta_{n+j(i)} \oplus G_1'(U_i \delta_{n+1+j(i)}) \oplus \cdots \oplus G_d'(U_i \delta_{n-d+j(i)}) = 0.$$ 

Summing them up, and we have

$$\bigoplus_{i=0}^{d} G_i'(U_j \delta_{n-i+k(j)}) = \bigoplus_{i=0}^{d} G_i'(V_j \delta_{n-i+k'(j)}) + \bigoplus_{i=0}^{d} G_i' \delta_{n-i}, \tag{4}$$

where $k(j) \in T_x$, $k'(l) \in T_z$, and $G_0' = E$. This process can be done for all parity-checks in $H^{(n)}$. The purpose is to determine $\delta_{n-i}$ of each position, when observing $\bigoplus_{j=1}^{\#T_e} V_j \delta_{n-i+k'(j)}$. Notice that the approach here is generic. When the parity-checks and linear approximations have special form, more efficient checking approach is feasible, see section 5.2.

There is no need that all $G_i = E, 1 \leq i \leq n$ as in linear distinguishing attack in large alphabets [YJM20], which is expected to have very high degree. For example, the degree of these special parity-checks with weight 4 of SNOW 3G is expected to be $O(2^{172})$.

To describe the effect of these parity-checks, we divide them into two sets. Let $H_I$ include those parity-checks whose coefficients are all $E$, while $H_{II}$ includes the rest. They are called type I and type II parities respectively, which play different roles in the iterative decoding phase.

### 3.3 Vectorial Iterative Decoding Algorithm

In this subsection, we consider how to extract information from a noisy sequence by vectorial iterative decoding algorithm. Firstly, we try to generalize original Algorithm B, then improve the iterative criterions.

Let $\#H^{(n)} = h$ denote the number of parity-checks with $r + 1$ taps at position (or clock) $n$. Let $e_1, \ldots, e_N$ denote the sequence of noises, and $z_1 \cdots z_N$ denote the deciphered sequence from key stream $z_1 \cdots z_N$ by $\bigoplus_{j \in \{1, \ldots, \#T_x\} \bigoplus_{i \in T_z} V_j \delta_{n+i+k'(j)}$. The initial priori distribution $P$ is the same for each $e_n$, which is derived by linear approximation. Let $p_{\zeta}^{(n)}(x) = \Pr[e_n = \zeta, \zeta \in \mathbb{F}_2^n]$ denote its density function, then the APP $p_{\zeta}^{(n)}(x)$ could be computed by Bayes’s formula.

$$p_{\zeta}^{(x)} = \Pr[e_n = \zeta | \text{when observed check values } (c_1, c_2, \ldots, c_b)] = p_{\zeta}^{(n)} \prod_{i \in H^{(n)}} \Pr[G_i' \delta_{n-l_i} = c_i \oplus E\zeta] = \bigoplus_{\eta} p_{\eta}^{(n)} \prod_{i \in H^{(n)}} \Pr[G_i' \delta_{n-l_i} = c_i \oplus E\eta] \tag{5}$$

We always assume $e_{n-l_i}$ are independent and all parity-checks are orthogonal. As $\zeta \in \mathbb{F}_2^n$ run over the alphabet, $\Pr[\sum_{i=1}^{r} G_i' \delta_{n-l_i} = c_i + E \zeta]$ can be calculated by convolution property and FWHT. Thus the nominator and denominator can be computed by Algorithm 2.

The vectorial iterative decoding algorithm is listed in Algorithm 3. The criterions which are used to break up the iterative loop and trigger the reset process are main factors affecting the convergence speed [CGD96, MG91].
Algorithm 2 Calculate the nominator

**Input:** priori p.d $p^{(n)}_\zeta$

1. Let priori probability distribution $p^{(n)} = (p_0, p_1, \cdots, p_{2^n-1})$.
2. For each parity-check $l \in H^{(n)}$
3. Calculate probability distribution $p(l)$ of $\sum_{i=1}^{\tau} G_i^l e_{n-l}$ by FWHT and convolution property.
4. Permute $p(l)x \leftarrow p(l)x \oplus \zeta$, $x \in \mathbb{F}_2^m$.
5. Multiply corresponding coordinate together of all these $p(l)$.

Algorithm 3 Vectorial iterative decoding

**Input:**
- The sequence $z'$ of length $N$ derived from key stream,
- The sequence of noises $e$ with initial p.d. $p$.
- The parity-checks set $H^d$ with $\tau + 1$ taps.

**Parameters:** Maximal rounds $R$, maximal iterations $T$ and minimal gap $G$ to infuse new noises.

1. Initialize the priori probability distribution sequence $\text{pri}$ of length $N$ all with the same initial probability distribution $p$.
2. Initialize the global empirical vector $E^{gb} = (E^{gb}_1, \cdots, E^{gb}_{2^m-1}) \leftarrow 0$.
3. For round $r = 1, 2, \cdots, R$ do
4. Initialize the round empirical vector $E^{rnd} = (E^{rnd}_1, \cdots, E^{rnd}_{2^m-1}) \leftarrow 0$.
5. Initialize the complement coin $\zeta \leftarrow 0$.
6. For iteration $i = 1, 2, \cdots, T$ do
7. Initialize a iteration empirical vector $E^{itr} = (E^{itr}_1, \cdots, E^{itr}_{2^m-1}) \leftarrow 0$.
8. For position $n = 1, 2, \cdots, N$ do
9. Compute $\text{app}$ from $\text{pri}$ by equation (5).
10. If $p^{(n)}_j > p^{(n)}_0$, then $E^{itr}_j \leftarrow E^{itr}_j + 1/N$, $j \in \{1, 2, \cdots, 2^m - 1\}$.
11. If $E^{itr} > E^{rnd}$, then $E^{rnd} \leftarrow E^{itr}$, $\text{pri} \leftarrow \text{app}$.
12. If $E^{itr} \leq E^{rnd}$ or $i = T$, then
13. If $E^{itr} = 0$, then return failed.
14. Else if $||E^{rnd} - E^{gb}|| < G$, then choose an very biased noise sequence $n$ of length $N$, reset $z' \leftarrow z' \oplus n$, break up current loop.
15. Else then $E^{gb} \leftarrow E^{rnd}$, select $\zeta$ such that $E^{itr}_{\text{init}(\zeta)} + E^{itr}_{\text{init}(\zeta)}$ is maximal, break up current loop.
16. If $\zeta \neq 0$, then complement all positions of $z'$ such that $p_\zeta > p_0$ with $\zeta$.
17. If $z'$ satisfies all parity-checks, then return success.
18. Reset a priori probability sequence $\text{pri}$ initial probability distribution $p$.
19. Terminate.
We try to optimize the criterions by experiments. Some phenomena are observed in scaled experiments when parity-checks are not so many. Firstly, if a threshold is raised to break up loop and reset as Algorithm B, it is easier to be triggered in the earlier rounds than the later rounds. Secondly, if a complement is performed very early without passing through enough iterations, it will pull the algorithm into a self-combination state very early and weaken the decoding efficiency. To improve this, two main criterions are proposed to break the loop and trigger the reset process.

**Criterion 1.** Passing through sufficient iterations before breaking up and resetting, which corresponding to line 7-11 and 14. More specifically, if new $\text{app}$ strengthens the empirical complement effect and iterations is less than maximal, then continue iteration by Bayes’s rule. Otherwise, select the complement coin which has potential largest empirical complement effect.

**Criterion 2.** When the empirical complement effect is weak from the previous round to current round, a sequence of very biased noises is infused in order to break the tie caused by self-composition property of LFSR. The noises’ SEI is required to be appropriate, neither very large to counteract the previous decoding work, nor very small to break the tie.

Criterion 1 is easy to understand. In order to avoid converging to self-combination state too early, we hope to correct errors as many as possible in each of the early rounds. Thus sufficiently iterations are needed before complementing. The idea behind criterion 2 is simple but novel. After many rounds, the complement would correct very few positions and weaken the decoding efficiency. To improve this, two main criterions are proposed to break up loop and reset as Algorithm B, it is easier to be triggered in the earlier rounds.

The complement in Algorithm 3 operates on derived sequence $z'$. The $n$-th position $z'_n$ is changed to $z'_n \oplus \zeta$ when the noise $e_n$ is determined to be $\zeta$ and the complement is performed. If $z'$ satisfies all parity-checks at the end, we just deduce that all $e_i = 0$. Then with the help of LFSR’s feedback polynomial, the initial state of LFSR can be recovered.

## 4 Cryptographic Properties and Experimental Results

### 4.1 Statistical Model

**Convergence Property**

It is necessary to figure out the convergence property when iteratively computing APP. Intuitively, we hope that APP $p^{*}_{\zeta(n)}$ increases when noise variable $e_n = \zeta$ and decreases when $e_n \neq \zeta$. Its expected value is computed as follows.

$$E_0[p^{*}_{\zeta}] = E[p^{*}_{\zeta}|e_n = \zeta]$$

$$= \sum_{c_1, \ldots, c_h} \frac{p^{(n)}_{\zeta} \prod_{l \in H(h)} \Pr[\sum_{i=1}^{r} G'_l e_{l_i} = c_l + E\zeta]}{\sum_{\zeta} p^{(n)}_{\zeta} \prod_{l \in H(h)} \Pr[\sum_{i=1}^{r} G'_l e_{l_i} = c_l + E\zeta]}$$

$$E_1[p^{*}_{\zeta}] = E[p^{*}_{\zeta}|e_n \neq \zeta]$$

$$= \sum_{\zeta \neq \zeta} \frac{p^{(n)}_{\zeta} \prod_{l \in H(h)} \Pr[\sum_{i=1}^{r} G'_l e_{l_i} = c_l + E\zeta]}{1 - p^{(j)}_{\zeta}}$$

And we conclude that $E[p^{*}] = p_{\zeta} E_0[p^{*}] + (1 - p_{\zeta}) E_1[p^{*}] = p_{\zeta}$. 

Zhou. Author et al.
Table 1: An example of increasing and decreasing ratio

<table>
<thead>
<tr>
<th>x</th>
<th>p_x</th>
<th>E_0/p_x</th>
<th>E_1/p_x</th>
<th>E_0/p</th>
<th>E_1/p</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>0.4500</td>
<td>1.02618712</td>
<td>0.97857418</td>
<td>0.097851734</td>
<td>0.987851734</td>
</tr>
<tr>
<td>1</td>
<td>0.2500</td>
<td>1.00117564</td>
<td>0.99960812</td>
<td>0.99313893</td>
<td>0.98837520</td>
</tr>
<tr>
<td>2</td>
<td>0.2000</td>
<td>1.02744428</td>
<td>0.9960812</td>
<td>0.99313893</td>
<td>0.98837520</td>
</tr>
<tr>
<td>3</td>
<td>0.2000</td>
<td>1.10462318</td>
<td>0.99960812</td>
<td>0.98837520</td>
<td>0.977212734</td>
</tr>
</tbody>
</table>

Example 1. Let the generator polynomial of LFSR $L(x) \in \mathbb{F}_2[x]$ with degree 16. We get the increasing and decreasing ratios in Table 1, when exploits 3 type I parity-checks with 3 taps. The second row is priori probability distribution $P$. $E_0/p$ and $E_1/p$ denote the increasing and decreasing ratio. Particularly, $E_0/p$ and $E_1/p$ denote the case only considering the number of holding parity-checks. Both cases meet our expectation.

Decoding Efficiency

In algorithm B, a threshold $N_{thr}$ is computed to promote the efficiency of the complements. It is determined by the intersection point of two shrunk normal distributions. In the multidimensional case, the intersection point becomes an intersection curve (surface). The threshold reflects the correcting ability of the first iteration in the binary case. Despite that we do not need such a threshold to promote efficiency in vectorial case, it still reflects the decoding efficiency from the first iteration. Thus we discuss how to estimate the correcting ability by measuring the volume of the intersection area in this subsection.

Let $N_{thr}^t$ denote this threshold corresponding to $\zeta$. Without loss of generality, we assume that the priori probability distribution $P$ of noise sequence $e_1 \cdots e_N$ s.t. $p_0 \geq p_1 \geq \cdots \geq p_{2^m−1} > 0$. Suppose that a random variable $X \sim P$, we require that the distribution of new random variable $G_i X$ still has 0 as the maximal value point 1. Obviously, it surely holds when $G_i$ is nonsingular. This requirement maybe reduce the number of available parity-checks, but it simplify the analysis for the effect of parity-checks.

Let $X_1, \cdots, X_r$ denote $\tau$ independent random variables all follows $P$. Let $Q$ denote the distribution of their linear combination $\sum_{i=1}^\tau G_i X_i$. Thus $Q$ still has 0 as its maximal value point, which could be deduced from the convolution property and Walsh-Hadamard transform. Particularly, if all $G_i = E$, $Q$ preserves the order of $P$, i.e., $q_0 \geq q_1 \geq \cdots \geq q_{2^m−1} > 0$.

The approach to calculate $N_{thr}^t$ is inspired by the fact $p_i^t$ is large when more check values appear to be $\zeta$. Let $q_0 = P\{\sum_{i=1}^\tau G_i e_{a−l} = c\}$ denote the probability that the $\tau$ taps sum to be $c$ for parity-check $l$. Obviously, $q_0$ depends on the individual parity-check. This phenomenon makes it very complicated to calculate the threshold $N_{thr}^t$. To simplify the calculation, we divide all parity-checks into two sets $\mathcal{H}_I$ and $\mathcal{H}_{II}$ according to its coefficients, then deal with them separately.

The set $\mathcal{H}_I$ includes all parity-checks whose coefficients are all identity. For this class, $q_0$ is obviously independent of parity-checks. Let $\#\mathcal{H}_I = h_I$, the probability the current noise $e = \zeta$ and $x_i$ check values equal $i, i \in \{0, \cdots, 2^m−1\}$ is as follows 2

$$p_c(q(x_0, \cdots, x_{2^m−1}, \zeta) = p_{x_0} \cdots p_{x_{2^m−1}} \prod_{i=0}^{2^m−1} q_{x_i | e = \zeta}^t.$$  

1Minimal value point is similar. We assume that $p_0$ is minimal instead.
2Actually, check values are vectors in $\mathbb{F}^2$, here we use integers $i$ to denote the same thing.
Figure 2: Example for the difference distribution

where \( x_{2m-1} = h - \sum_{i=0}^{2m-2} x_i \).

Obviously, \( x = (x_0, \cdots, x_{2m-1}) \) follows multinomial distribution \( \text{Multi}(h, q_\zeta) \) with parameter \( q_\zeta = (q_\zeta, \cdots, q_{2m-1}) \). Its density function are denoted by \( q(x, \zeta) \). For convenience, we introduce notations

\[
q^\pi = \prod_{i=0}^{2m-1} q_{\pi_{i\oplus\zeta}}, \quad h = \frac{h_1!}{x_0! \cdots x_{2m-1}!}.
\]

Let \( \mathcal{A}(\zeta) \) be a subset of all possible \( x \). Once we complement those noises with \( \zeta \neq 0 \) when the vectors in \( \mathcal{A}(\zeta) \) are observed, the expected number of correctly complemented noises and erroneously complemented noises are respectively

\[
N \times W(P, \mathcal{A}(\zeta), \zeta) = N \sum_{x \in \mathcal{A}(\zeta)} p_\zeta q(x, \zeta), \quad N \times W(P, \mathcal{A}(\zeta), 0) = N \sum_{x \in \mathcal{A}(\zeta)} p_0 q(x, 0), \quad (7)
\]

where \( N \) denote the length of data. All the other cases of complements are neutral. Thereby, the number of actual corrected positions is the difference

\[
N \times I(P, \mathcal{A}(\zeta), \zeta, 0) = N \times W(P, \mathcal{A}(\zeta), \zeta) - N \times W(P, \mathcal{A}(\zeta), 0). \quad (8)
\]

Given \( P \) and \( H_1 \), if we can find a set \( \mathcal{A}(\zeta) \) maximizing \( I(P, \mathcal{A}(\zeta), \zeta, 0) \), then the expected number of actual corrected positions of each complement should be maximized. Firstly, we observe that the means of the two multinomial distributions are \( h_1 q_\zeta \) and \( h_1 q_0 \) respectively. Therefore, similar as the binomial case, there is a set \( \mathcal{A}(\zeta) \) of \( x \) in which \( I(P, \mathcal{A}(\zeta), \zeta, 0) \) takes non-negative value.

Since given \( x \), \( I(P, \mathcal{A}(\zeta), \zeta, 0) \) and \( p_\zeta^x - p_0^x \) have the same sign, it is equivalent to find \( \mathcal{A}(\zeta) \) such that \( p_\zeta^x - p_0^x > 0 \) for each \( x \in \mathcal{A}(\zeta) \) , that is to determine the region \( \mathcal{A}(\zeta) \) such that

\[
\delta(\zeta, 0) = p_\zeta q(x, \zeta) - p_0 q(x, 0) > 0, \quad x \in \mathcal{A}(\zeta). \quad (9)
\]

**Example 2.** Let initial distribution \( P \) and LFSR be the same as in Example 1, and \( h_1 = 15 \). The difference \( \delta(\zeta, 0) \) is illustrated in Fig. 2. The non-negative and the negative area are separated. The size of circle represents the relative absolute value of the difference \( \delta(\zeta, 0) \).
Table 2: Direct computation and normal approximation for $I(p, A(1), 1, 0)$

<table>
<thead>
<tr>
<th>Number of equations $h_I$</th>
<th>40</th>
<th>80</th>
<th>200</th>
<th>400</th>
</tr>
</thead>
<tbody>
<tr>
<td>Direct computation</td>
<td>0.0686</td>
<td>0.1138</td>
<td>0.1835</td>
<td>0.2266</td>
</tr>
<tr>
<td>Normal approximation</td>
<td>0.0707</td>
<td>0.1148</td>
<td>0.1841</td>
<td>0.2267</td>
</tr>
</tbody>
</table>

When $h$ is small, it is feasible to evaluate $N_{\zeta}^{thr}$ by exhaustively searching. The threshold $N_{\zeta}^{thr}$ can be determined by

$$N_{\zeta}^{thr} = N\left( \sum_{x \in A(\zeta)} \sum_{\eta \in F_2^{m}} p_\eta q(x, \eta) \right).$$

The time complexity is about $O(2^m \left( h_I + 2^m \right))$.

When $h_I$ is large and $q$ is not near the boundary of the parameter space, multivariate normal distribution approximation is suitable. Multi($h_I, q$) could be approximated by $N(\mu, \Sigma)$ with density function

$$\frac{1}{\sqrt{(2\pi)^m |\Sigma|}} \exp \left( -\frac{1}{2} (x - \mu)^T \Sigma^{-1}(x - \mu) \right),$$

where superscript $T$ denotes transposition, mean vector $\mu$ and covariance matrix $\Sigma$ are determined by Multi($h_I, q$). Therefore, the area $A(\zeta)$ maximizing the multiple integral

$$I(P, A(\zeta), \zeta, 0) \approx \int_{A(\zeta)} \left( p_\zeta N(\mu_\zeta, \Sigma_\zeta) - p_0 N(\mu_0, \Sigma_0) \right) dx$$

should be part of a hypercube with dimension $2^m - 2$ that restricted by the $2^m - 1$ coordinate plane and two surfaces

$$\Omega_1: \sum_{i=1}^{2^m - 2} x_i = h_I,$$

$$\Omega_2: \frac{1}{2} ((x - \mu_0)^T \Sigma_0^{-1}(x - \mu_0)) - \frac{1}{2} ((x - \mu_\zeta)^T \Sigma_\zeta^{-1}(x - \mu_\zeta)) - \ln \frac{p_0}{p_\zeta} = 0.$$

Notice that $\Omega_2$ is a quadratic form in the real field, the multiple integral (11) can be computed by repeated integral. Once $A(\zeta)$ is determined, the threshold can be calculated by volume integral

$$N \sum_{\eta \in F_2^m} \int_{A(\eta)} N(\mu_\eta, \Sigma_\eta) dx.$$

Example 3. Let the probability distribution $P$ and LFSR be the same as in Example 1. To illustrate this multivariate normal approximation, $I(P, A(1), 1, 0)$ is computed by two methods and depicted in Table 2. In order to simplify the integral, we could even slightly adequate the boundary of $A$ without fluctuating the result much.

When the parity-checks stem from $H_{II} = H \setminus H_I$, $q_c$ depends on individual parity-check. Thus when the probability value peak is $q_0$, we introduce a symmetric multinomial distribution $Q'$ to simulate the influences of type II parity-checks, which parameter is

$$q'_0 = q_0, q'_1 = \cdots = q'_{2^m - 1} = \frac{1 - q_0}{2^m - 1}.$$
Table 3: Theoretical and empirical value of $N^\text{thr}_\zeta / N$

<table>
<thead>
<tr>
<th>No. of parities $(h_I, h_{II})$</th>
<th>$\zeta$</th>
<th>theoretical</th>
<th>empirical $N = 2^{19}$</th>
<th>empirical $N = 2^{20}$</th>
<th>empirical $N = 2^{21}$</th>
</tr>
</thead>
<tbody>
<tr>
<td>(36,0)</td>
<td>1</td>
<td>0.277133</td>
<td>0.227242</td>
<td>0.250517</td>
<td>0.264012</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>0.253926</td>
<td>0.242359</td>
<td>0.246835</td>
<td>0.249339</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>0.200412</td>
<td>0.164480</td>
<td>0.181245</td>
<td>0.190250</td>
</tr>
<tr>
<td>(18,18)</td>
<td>1</td>
<td>0.297959</td>
<td>0.251286</td>
<td>0.270056</td>
<td>0.279394</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>0.260769</td>
<td>0.220915</td>
<td>0.238914</td>
<td>0.248543</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>0.167968</td>
<td>0.125576</td>
<td>0.144096</td>
<td>0.154273</td>
</tr>
<tr>
<td>(0,138)</td>
<td>1</td>
<td>0.376058</td>
<td>0.360392</td>
<td>0.364783</td>
<td>0.368026</td>
</tr>
<tr>
<td></td>
<td>2</td>
<td>0.325561</td>
<td>0.321800</td>
<td>0.332389</td>
<td>0.338674</td>
</tr>
<tr>
<td></td>
<td>3</td>
<td>0.221771</td>
<td>0.198662</td>
<td>0.213513</td>
<td>0.221388</td>
</tr>
</tbody>
</table>

Example 4. To verify the validity of these approximations, with the same $P$ and LFSR as in Example 1, we compute the theoretical ratio of $N^\text{thr}_\zeta / N$ and the empirical ratio by the ratio where $p_\zeta > p_0$. Table 3 depicts that our estimations are very precise.

4.2 Information Theory Properties

In this subsection, we discuss some properties from the point view of information theory. Suppose the noises are independent and the parity-checks are linear independent, the relative entropy between Multi$(h, q_0)$ with density function $q(x)$ and Multi$(h, (2^{-m}, \cdots, 2^{-m}))$ with density function $u(x)$ is

$$D(q \parallel u) = H(q, u) - H(q) = h \sum_{i=0}^{2^m-1} q_i \log \frac{q_i}{2^{-m}} = h(m - H(q_0)). \quad (15)$$

That is the relative entropy is the number of parity-checks times the SEI of probability distribution $Q$.

Secondly, we hope that the right corrected positions are as many as possible in the complement process. Now we think about the sum of relative entropy between Multi$(h, q_c)$ and Multi$(h, q_0)$ for all $c \neq 0$, and we have

**Proposition 1.** Let $q_c(x)$ and $q_0(x)$ be density functions of Multi$(h, q_c)$ and Multi$(h, q_0)$ respectively, then

$$\sum_{c \neq 0} D(q_c(x) \parallel q_0(x)) = -h \log \prod_{i=0}^{2^m-1} q_i - h2^m H(q_0).$$
Proof.

\[
\sum_{c \neq 0} D(q_c(x) \parallel q_0(x)) = \sum_{c \neq 0} h(\sum_{i=0}^{2^m-1} q_i \log q_i - \sum_{i=0}^{2^m-1} q_i \log q_i^c).
\]

\[
= -h(2^m - 1)H(q_0) - h \sum_{i=0}^{2^m-1} q_i \log q_i.
\]

\[
= -h \log \prod_{i=0}^{2^m-1} q_i - h2^m H(q_0).
\]

This tells us when the probability distribution of noises approaches uniform distribution, the total relative entropy converges to 0.

4.3 Complexity Analysis

On one hand, given the SEI \( \Delta(p) \), the code rate \( k/N < \Delta(p)/(2 \ln(2)) \) to transmit \( k \) bits information through a SC channel by Shannon’s Theorem. On the other hand, the number of parity-checks \( h \) influences the decoding complexity. We focus on the property of the first iteration in the first round, which seems to be the critical part by previous section, and discuss how to deduce some theoretical bounds for \( h \) as well as key stream length \( N \).

A Bound Derived from Decoding Codes

Similarly as Proposition 1 in [CS91], In order to perform an error corrected iterative decoding, the lower bounds of \( h \) should satisfy that there exists at least a \( \zeta \) such that \( p^{*}_{\zeta} > p^{*}_0 \). It is summarized as follows.

Proposition 2. If iterative decoding is feasible, then there is at least one \( \zeta \in \{1, 2, \cdots, 2^m - 1\} \) such that \( p^{*}\zeta q(x, \zeta)/(p^{*}_0 q(x, 0)) > 1 \). Particularly, when \( P \), \( Q \) and \( Q' \) are multinomial probability distributions as before, then \( \zeta = 2^m - 1 \) and

\[
\frac{p^{*}_{\zeta}}{p^{*}_0} > \left( \frac{q^{*}_{\zeta}}{q^{*}_0} \right)^{b_{1i}} \left( \frac{q^{*}_{\zeta}}{q^{*}_0} \right)^{b_{1i}}.
\]

Proof. Since if \( p^{*}\zeta q(x, \zeta)/(p^{*}_0 q(x, 0)) \leq 1 \) holds for all \( \zeta \), then \( p^{*}_i \) converges to 0 or becomes ambiguous during the iterations, i.e., \( p^{*}_0 = p^{*}_1 \) is one of the largest. The decoding algorithm won’t work.

Particularly, when the probability values of \( P \) and \( Q \) are in order as stated before, and all values of parity-checks are \( \zeta \), obviously we have

\[
\frac{p^{*}_{\zeta} q^{*}_{\zeta}}{p^{*}_0 q^{*}_0} \leq \frac{p^{*}_{\zeta} q^{*}_{\zeta} b_{1i} b_{1i}}{p^{*}_0 q^{*}_{0} b_{1i} b_{1i}}.
\]

Remark 2. Though the ratio \( \eta(\zeta, 0) \) has large value when all check values are \( \zeta \), The lower bound for \( h \) given in Proposition 2 may be loose, as the probability that all check values are \( \zeta \) is small.
A lower bound for $N$ could be derived through Proposition 2. For example, when generator polynomial $L(x) \in \mathbb{F}_{2^m}[x]$, the number of parity-checks $h$ and the key stream length $N$ shall satisfy that $\binom{N}{2^{m-1}} \approx h 2^{k}$. Hence, by Proposition 2, the number of type I and II parity-checks for $P$ is mainly benefits from the increased SEI. More specifically, according to Theorem 1, there are two important probability distributions. Since when $\Delta(e) = 2^\gamma$, it is expected that there is a probability value around $2^{-m} \pm 2^{-\frac{m+\gamma}{2}}$ in practice [YJM20], the distributions $P$ and $P'$ in Table 4 is very likely to appear, where $\varepsilon$ denotes $(1 - 2^{-m} - 2^{-\frac{m+\gamma}{2}})/(2^m - 1) - 2^{-m}$.

By Taylor’s formula, we have

$$
p_i = 1 - \frac{\tau}{2^m - 1}, p'_i = 1 - \frac{2^m(\tau + 1)}{2^m - 1},
$$

Furthermore, by the convolution property, when each parity-check has $\tau + 1, \tau \geq 2$ taps, we have

$$
q_i = 1 - \frac{2^m(\tau + 1)}{2^m - 1} = 1 - \frac{2^m(\tau + 1)}{2^m - 1}.
$$

Hence, by Proposition 2, the number of type I and II parity-checks for $P$ are

$$
1 - \frac{2^m(\tau + 1)}{2^m - 1} \geq \left(1 - 2^{-2^m(\tau + 1)(\tau - 1) + 2}\right) h_i \Rightarrow h_i \geq 2^{-2^m(\tau + 1)(\tau - 1) + 2},
$$

where $2^m(\tau - 1) \approx 1$.

For the case of $P'$, the general term formula of distributions convolution could be deduced by its recursion formula, i.e.,

$$
q'_i = 2^{-m} + \frac{2^m(\tau - 1)}{2^m - 1} 2^{-2^m(\tau - 1) + 1} 2^{-\frac{m+\gamma}{\tau}}, q'_i = 2^{-m} - \frac{2^m(\tau - 1)}{2^m - 1} 2^{-\frac{m+\gamma}{\tau}}.
$$

Thus we have

$$
q'_i \approx 1 - \frac{2^m(\tau + 1)}{2^m - 1} 2^{-\frac{m+\gamma}{\tau}},
$$

which means

$$
1 - \frac{2^m}{2^m - 1} 2^{-\frac{m+\gamma}{\tau}} \geq \left(1 - \frac{2^m(\tau + 1)}{2^m - 1} 2^{-\frac{m+\gamma}{\tau}}\right) h \geq \left(\frac{2^m - 1}{2^m}\right)^{\tau - 1} 2^{-\frac{m+\gamma}{\tau}}.
$$

Notice that type I and II parities are not distinguished in the case of $P'$.

We expected that the practical bound is between those deduced by $P$ and $P'$. The FCA mainly benefits from the increased SEI. More specifically, according to Theorem 1, there are $2^m - 1$ binary linear approximations contributing to the SEI of linear approximation with dimension $m$.

Notice that there are another distributions, e.g., $P''$ with $p''_i = 2^{-m} - 2^{-\frac{m+\gamma}{2}}$, while the other value point are all the same. This case is similar with $P'$ except that $0$ is the minimal value point.
A Bound Derived from the Practical Corrected Errors

In this part, we discuss how to deduce a bound from the number of expected positions with $p_{c}^{\ast} > p_0$, $\zeta \neq 0$.

Let us consider the sets $\mathcal{A}(i), i \in \{1, 2, \cdots, 2^{m} - 1\}$ for multinomial distributions. Since $\mathcal{A}(i)$ may intersect with each other, the way of computing threshold in section 4.1 can’t be directly applied. Thereby, we introduce some new sets: $\mathcal{A}'(i) = \mathcal{A}(i) - \mathcal{A}(i) \cap \bigcup_{j=1}^{i-1} \mathcal{A}(j)$, That is $\mathcal{A}(i)$ excluding all elements that are included in previous sets $\mathcal{A}(i), i \in \{1, 2, \cdots, i\}$. Let $M_{i}'$ denote the summation of probability values over set $\mathcal{A}'(i)$, more specifically,

$$\sum_{i=1}^{2^{m} - 1} p_{\zeta} \sum_{x \in \mathcal{A}'(\zeta)} q(x, \zeta). \quad (19)$$

It is reasonable to require that $\sum_{i=1}^{2^{m} - 1} M_{i}' > 1$ after the first iteration. Then the succeeding iterations may trigger more positions with $p_{c}^{\ast} > p_0^{\ast}$. This phenomenon may be the main advantage that soft decision decoding algorithms have.

Summing up the probability values in multinomial distributions is inconvenient. Though multivariate normal distribution approximation could also be used as before when $h$ is large, the integral may not be easy to evaluate in practice, as the integral area $\mathcal{A}'(\zeta)$ is very complicated. Since symmetric distribution $Q'$ simulates the iterative process very well, we could deduce boundaries for $\mathcal{A}'(\zeta)$ using Multi($h, q'$). The following results shows how to estimate $M_{i}'$ in this case.

**Proposition 3.** For multinomial probability distribution Multi($h, q'$), we have

$$M_{i}' = \sum_{i=1}^{h} \binom{h}{i}(1 - \sum_{i=0}^{\zeta} q_{i}^{\prime}_\zeta)^{i-1} \sum_{(x_0, \cdots, x_\zeta) \in \mathcal{B}(\zeta)} \left( q_{i}^{\prime}_\zeta \right)^{l}_{i=0} \prod_{i=0}^{l} q_{i}^{\prime}_\zeta,$$

where $\mathcal{B}(\zeta)$ is constrained by $\sum_{i=0}^{\zeta} x_i = l$, $x_\zeta - x_0 \geq h_b$ and $x_i - x_0 \leq h_b$, $1 \leq i < \zeta$.

Particularly, when $\sum_{i=0}^{\zeta} q_{i}^{\prime}_\zeta$ is small and $h q_{1}^{\prime} \leq h_b$, the expected number of positions with $p_{c}^{\ast} > p_0^{\ast}$ in the first iteration are dominated by those small $l$.

**Proof.** Since $q_{1}^{\prime} = \cdots = q_{2^{m} - 1}$, we have

$$M_{i}' = \sum_{x \in \mathcal{A}'(\zeta)} q_{x}^{\zeta} = \sum_{i=1}^{h} \binom{h}{i}(1 - \sum_{i=0}^{\zeta} q_{i}^{\prime}_\zeta)^{i-1} \sum_{(x_0, \cdots, x_\zeta) \in \mathcal{B}(\zeta)} \left( q_{i}^{\prime}_\zeta \right)^{l}_{i=0} \prod_{i=0}^{l} q_{i}^{\prime}_\zeta.$$

By Proposition 2, we deduce that there is a minimal positive integer $h_b$ such that $\delta(\zeta, 0) > 0$ when $x_\zeta - x_0 \leq h_b$. Furthermore, $x_i - x_0 < h_b$ should holds for all $0 < i < \zeta$ to exclude the points in $\mathcal{A}'(i)$. Therefore, when $p_{c}^{\ast} > p_0^{\ast}$, $(x_0, \cdots, x_\zeta) \in \mathcal{A}'(\zeta)$ must satisfy that

$$\begin{cases} x_i \geq 0, & 0 \leq i \leq \zeta, \\ x_i - x_0 < h_b, & 0 < i < \zeta, \\ x_\zeta - x_0 \geq h_b, \\ x_0 + \cdots + x_\zeta < h. \end{cases}$$

When $h$ is not small and $\sum_{i=0}^{\zeta} q_{i}^{\prime}_\zeta$ is not high, multidimensional distribution Multi($h, q'$) could be approximated by $\zeta + 1$ independent Poisson distributions with means $\lambda_{i}^{\prime} = h q_{i}^{\prime}_\zeta$, i.e.,

$$\Pr(X = x) \approx \sum_{\mathcal{A}'(\zeta)} \prod_{i=0}^{\zeta} \frac{\lambda_{i}^{\prime}}{x_i!} e^{-\lambda_{i}^{\prime}} = \frac{\lambda_{0}^{\zeta}}{x_0!} e^{-\lambda_{0}} \frac{\lambda_{0}^{x_0}}{x_0!} e^{-\lambda_{0}} \cdots \frac{\lambda_{\zeta}^{x_\zeta-1}}{x_\zeta-1!} e^{-\lambda_{\zeta}}. \quad (20)$$
As $\lambda_i \leq h_b$, the maximal value of $\Pr(X = x)$ is when $\sum_{i=0}^{\zeta} x_i$ is small, i.e., when $l$ is small.

Proposition 3 gives us a hint that the value corresponding small $l$ dominate $M'_1$. When $\zeta$ is not very large, $M'_1$ could be approximated by partial summation for small $l$ close to the boundary. Obviously, $M'_{i \neq 0}$ are monotone non-increasing sequence.

When $\zeta = 1$, there is another elegant way to estimate $M'_1$ by Skellam distribution. Let $Y_0 \sim \text{Pois}(\lambda_1)$ and $Y_1 \sim \text{Pois}(\lambda_0)$, we know that their difference $K = Y_1 - Y_0$ follows Skellam distribution with following probability density function.

$$p(k, \lambda_1, \lambda_0) = e^{-\lambda_1 - \lambda_0} \frac{(\lambda_1/\lambda_0)^{k/2}}{I_{|k|}(2\sqrt{\lambda_1 \lambda_0})},$$

where $I_{|k|}$ is the modified Bessel function of the first kind. Obviously, $M'_1 = N p_k \Pr(K > h_b)$ since a boundary line is $x_1 \geq x_0 + h_b$ by Proposition 3.

### On Sparse Check Equations

Since sparse parity-checks have large advantages while checking parity, we are interested in these parity-checks with $\tau = 1$ or 2. In this section, we give some miscellaneous observations about them.

Let $x_t = (x_{t+c_1}, x_{t+c_2}, \ldots, x_{t+c_m})$, $c_1 < \cdots < c_m$ denotes the output at time $t$ of LFSR with generator polynomial $L(x) \in M_m(F_2)[x]$. Each coordinate sequence is a $m$-sequence $x_1 x_2 \cdots$ left shifting $c_i$ times, and its minimal polynomial $f(x) \in F_2[x]$ has degree $k$. Particularly, when shift vector $(c_1, c_2, \cdots, c_m)$ satisfies special condition, it becomes an LFSR over extension field $F_{2^m}$ [GX94].

Though parity-checks with $\tau + 1 = 2$ taps have very large advantages, unfortunately, the existence of them is a problem by the following direct observations.

**Proposition 4.** Let $x_t = (x_{t+c_1}, x_{t+c_2}, \cdots, x_{t+c_m})$ be as stated above, we have

- If $c_m - c_1 + m - 1 < k$, then there is no parity-check with $\tau = 1$.
- Given two parity-checks with $\tau = 1$, $G x_t + E x_{t+d_1} = 0$, $G' x_t + E x_{t+d_2} = 0$, if $d_1 = d_2$ and $x_t$ run over all values in $F_2^m \setminus \{0\}$, then $G = G'$. If $d_1 \neq d_2$, then $\gcd(d_1, d_2) > k - m$.

**Proof.** 1. Let $G x_t + E x_{t+d} = 0$ be a parity-check. Since $i$-th row of $A$ and $E$ forms a check polynomial $f_i$ with nonzero constant for $x_t$, then $f_i f_s$. As $G$ is nonsingular, there must be two different check polynomials $f_i(x)$ and $f_j(x)$. That means $f_1 + f_2$ also forms a check polynomial, but $c_m - c_1 + m - 1 < k$ means a polynomial with degree less than $k$ could be deduced, which is impossible.

2. When $d_1 = d_2$, it is deduced that $(G + G') x_t = 0$ for all $x_t$. When $x_t$ run over all values in $F_2^m \setminus \{0\}$, then we have $G = G'$.

When $d_1 < d_2$, we could deduce another linearly dependent parity-check

$$G'(G^{-1} x_t + E x_{t+d_2-d_1}) = 0.$$

Therefore, according to Euclid long division algorithm, there is a $G^*$ which satisfies

$$G^* x_t + E x_{t+gcd(d_1, d_2)} = 0.$$

Since there are $k$ information bit of LFSR, then $gcd(d_1, d_2) \geq k - m$. \qed
These observations imply that parity-checks with \( \tau = 1 \) may be rare, but it doesn’t mean none, even though the key stream length needed may be large. For example, when all \( (x_{t+c_1}, x_{t+c_2}, \cdots, x_{t+c_m}) \) are only in a subspace of \( \mathbb{F}_2^m \), and \( c_m - c_1 + m - 1 \) is large. Once a parity-check is found, more could be constructed by sliding and adding together.

Moreover, if a parity-check satisfies sequence \( x_i \), then its characteristic polynomial \( F_n(x) \in \mathbb{F}_2[x] \) has \( f(x) \) as a factor. Since \( G_n = G, G_1 = \cdots = G_{n-1} = 0 \), then \( F_n(x) = \text{det}(Ex^n + G) \), the number of choices for matrix \( G \) and \( n = (N/m - 1)|GL_m(\mathbb{F}_2)| \). Let \( S = \{ F_n(x) : 1 \leq nm \leq N \} \) denote all possible characteristic polynomials. For convenience, we introduce a map sending \( F_n(x) \in S_G \) to \( \mathbb{F}_2[x] \).

\[
\phi : S_G \to \mathbb{F}_2[x] \\
F_n(x) = \text{det}(Ex^n + G) \to F(x) = \text{det}(Ex + G).
\]

Since \( F(x) \) is the characteristic polynomial of invertible matrix \( G \), the number of different \( F(x) \) is \( 2^{m-1} \). Suppose that \( F(x) = f_1^{m_1} \cdots f_n^{m_n} \), where all \( f_i \) are distinct irreducible polynomials of degree \( d_i \), it has been proved that the number of \( G \) with given \( F(x) \) is \( \theta(F(x)) \) [Ger61], i.e.,

\[
\theta(F(x)) = 2^{m-1} \frac{\prod_{i=1}^{m} (1 - 2^{-i}) \prod_{i=1}^{m} (1 - 2^{-j d_i})}{\prod_{i=1}^{m} (1 - 2^{-i d_i})}.
\]

We also know that \( F_n(x) = F_{n+i}(x)^2 \) for some \( i > 0 \) when \( n_1 \) and \( n_2 \) are in the same 2-cyclotomic coset \( C \) modulo \( \text{ord}(f) = 2^k - 1 \). And the size of set \( F = \{ F_n(x) : 1 < nm < N \} \) is bounded by \( N/(km) < \#F \leq \sum_{d|k} \mu(d) \sum_{i=1}^{k/d} 2^i \), where \( \mu(\cdot) \) is Möbius function.

For the case \( \tau \geq 2 \), there are about \((N/m - 1)|GL_m(\mathbb{F}_2)| (2^{m-1} - 1) \) choices for the two coefficients and \( n \). An upper bound of \( \#S \) is the number of conjugacy classes of \( T \) in \( GL_m(\mathbb{F}_2) \), which is roughly about \( 2^{m-1} - \sum_{l=1}^{[nm/3]} 2^l \). We believe it is much more than \( (2^m - 1)^2 \) when \( L(x) \in \mathbb{F}_2^m \).

The Case for \( m = 1 \)

Regardless of the differences in criterions, the original FCA proposed by Meier et. al. can be treated as a special case of new FCA with dimension \( m = 1 \). The coefficient matrices of LFSR degenerates to scalar elements in \( \mathbb{F}_2 \). Therefore, the commutative condition for coefficient matrices of parity-checks is no need to be considered. The multidimensional linear approximations degenerates to binary linear approximation. Since the multinomial distribution degenerates to binomial distribution, \( \zeta \) must be 1. The bound derived from Proposition 2 is the same as in [CS91]. Estimating \( M_1^t \) also becomes simple.

Small Scale Experiments

We perform a scaled experiment to verify the vectorial iterative decoding algorithm in this subsection. The experiment settings are as follows. The generator polynomial of LFSR is \( g(x) = x^{16} + x^{15} + x + \alpha \in \mathbb{F}_2[x] \), where \( \alpha \) is the primitive element of \( \mathbb{F}_2^2 \). The output of LFSR at time \( t \) is \( x_t \). The noise stems from a SC channel instead of nonlinear part of a stream cipher. The target is recovering LFSR output sequence \( x_1x_2 \cdots x_N \) from noisy sequence \( z_1z_2 \cdots z_N = (x_1x_2 \cdots x_N) \oplus (e_1e_2 \cdots e_N) \).

We tweak the parameters such as channel capacity, the number of parity-checks and the infused noise to verify the word-error ratio(WER) after iterating a number of rounds. Specifically, the density functions of 2 priori distributions \( P_1 \) and \( P_2 \) are \((0.45, 0.25, 0.2, 0.1) \) and \((0.33, 0.25, 0.22, 0.20) \) respectively. The length of datat is \( N = 2^{19} \) or \( 2^{21} \) key stream words. The number of parity-checks with \( \tau = 2 \) are \( h = 9, h_1 = 36 \) or \( h_1 = 36 \). The results of experiment are illustrated in 3. For example, the curve \( (1, 1, 2, 9, 19) \) denotes the result derived by parameters \( P_1, h = 9, N = 2^{19} \) with Criterions 1 and 2. The curve
(2, thr, -, 36II, 19) denotes the result derived by parameters $P_2$, $h_{II} = 36$, $N = 2^{19}$ with threshold criterion like Algorithm B.

Some observations could be induced from Figure 3. Firstly, comparing the curve $(1, 1, 2, 9, 19)$ with $(1, 1, 2, 36I, 19)$, we see that the convergence speed increases with the number of parity-checks when channel capacity is fixed. Secondly, infusing new noise indeed increases the convergence speed. Thirdly, Criterion 1 increases the convergence speed. Notice that $(2, 1, 2, 36I, 19)$ seems worse than $(2, 1, 2, 36II, 19)$. The reason is that the length of key stream $N = 2^{19}$ is not sufficiently large comparing with the degrees. Therefore, the average feasible parity-checks for both the head and tail segments of the key stream in $(2, 1, 2, 36I, 19)$ are less than in $(2, 1, 2, 36II, 19)$.

5 Application to Grain-128a

In this section, we apply our new techniques to stream cipher Grain-128a. We assume the cryptanalysis is under the known-plaintext scenario. Since the output is directly used as key stream and the plaintext never participates in updating internal states, this assumption is reasonable for Grain-128a.

5.1 A Brief Description of Grain-128a

Grain-128a includes a 128-bit LFSR cascaded with a 128-bit NFSR. Let $s^{(t)} = (s_t, s_{t+1}, \ldots, s_{t+127})$ and $b^{(t)} = (b_t, b_{t+1}, \ldots, b_{t+127})$ denote their internal states at time $t$. The output $y_t$ of the pre-output function at time $t$ is represented by

$$y_t = h(s^{(t)}, b^{(t)}) \oplus s_{t+93} \oplus b_{t+2} \oplus b_{t+15} \oplus b_{t+36} \oplus b_{t+45} \oplus b_{t+64} \oplus b_{t+73} \oplus b_{t+89},$$

where $h(s^{(t)}, b^{(t)})$ is defined as

$$h(s^{(t)}, b^{(t)}) = h(b_{t+12}, s_{t+8}, s_{t+13}, s_{t+20}, b_{t+95}, s_{t+42}, s_{t+60}, s_{t+79}, s_{t+94})$$
$$= b_{t+12}s_{t+8} \oplus s_{t+13}s_{t+20} \oplus b_{t+95}s_{t+42} \oplus s_{t+40}s_{t+79} \oplus b_{t+12}b_{t+95}s_{t+94}. $$
The feedback bits of LFSR and NFSR are computed by

\[
s_{t+128} = s_t \oplus s_{t+7} \oplus s_{t+38} \oplus s_{t+70} \oplus s_{t+81} \oplus s_{t+96},
\]

\[
b_{t+128} = s_t \oplus b_t \oplus b_{t+26} \oplus b_{t+56} \oplus b_{t+91} \oplus b_{t+96} \oplus
\]

\[
    b_t + 3b_t + 67 \oplus b_{t+11}b_{t+13} \oplus b_{t+17}b_{t+18} \oplus b_{t+27}b_t + 59 \oplus
\]

\[
b_{t+40}b_{t+48} \oplus b_{t+61}b_{t+65} \oplus b_{t+68}b_{t+84} \oplus
\]

\[
b_{t+22}b_{t+24}b_{t+25} \oplus b_{t+70}b_{t+78}b_t + 82 \oplus b_{t+88}b_{t+92}b_t + 93b_t + 95.
\]

Key stream bit \( z_t = y_t \) in the stream cipher mode, while \( z_t = y_{2w+t} \) in the authenticated mode, where \( w \) is the tag size. The overall structure of Grain-128a is depicted in Fig. 4.

### 5.2 Constructing Multidimensional Linear Approximations and Checking Parity

In [TIM+18], the authors proposed a family of linear approximations of Grain-128a by piling up different clocks to eliminate the linear terms of the NFSR, which forms are

\[
\oplus_{i \in T_z} y_{t+i} \approx \oplus_{i \in T_z} s_{t+i+93} \oplus \oplus_{j \in A} s_{t+i+j} \oplus \oplus_{i \in T_z} (\Lambda_i[1 - 3], s_{t+i+8}, s_{t+i+13}, s_{t+i+20}) \oplus (\Lambda_i[5 - 8], s_{t+i+42}, s_{t+i+60}, s_{t+i+79}, s_{t+i+94}),
\]

(21)

where \( A = \{2, 15, 36, 45, 64, 73, 89\}, T_z = \{0, 26, 56, 91, 96, 128\}, \Lambda_i \) is a 9-bit binary linear mask, two bits \( \Lambda_i[0, 4] \) are fixed.

According to [TIM+18], an assignment of \( \Lambda_i[1 - 3] \) and \( \Lambda_i[5 - 8] \) will completely determine the correlation of \( h \) function, when \( \Lambda_i[0, 4] \) is fixed. For a specific \( i \in T_z \), there are only 64 possible \( \Lambda_i[0, 4], i \in A \) such that the correlation of Eq. (21) is nonzero. Hence, the linear correlation value of (21) can be deduced by summing up all these 64 \( \Lambda_i[0, 4], i \in T_z \). Meanwhile, there are \( 2^6 \) values of \( \Lambda_i[1 - 3, 5 - 8] \) of a specific \( i \in T_z \) with the correlation of \( h \) function is nonzero. For example, when \( \Lambda_i[1 - 3, 5 - 8] = 00000000, \forall i \in T_z \), the correlation of (21) is about \( \pm 2^{-57.0454} \). For more details of these linear approximations, we refer to [TIM+18].

In this paper, we reuse these linear approximations but in a new way by bundling them up. Firstly, we choose 42 linear approximations which \( \Lambda_i[1 - 3, 5 - 8], i \in T_z \) has form

\[
(\Lambda_0[1 - 3, 5 - 8], \Lambda_20[1 - 3, 5 - 8], \cdots, \Lambda_{128}[1 - 3, 5 - 8]) = (0, \cdots, 0, 1, 0, \cdots, 0),
\]

i.e., \( \Lambda_i[1 - 3, 5 - 8], i \in T_z \) as a group of standard basis. Then a linear approximation with dimension \( 9 \leq m \leq 42 \) can be established as follows

\[
E(x_t + u_t) + Ey_t = e_t,
\]

(22)
where $E$ is an $m \times m$ identity matrix in $\mathbb{F}_2$, $e_t$ is noise vector, and

$$x_t = \left( \cdots, s_{t+i+8}, s_{t+i+13}, s_{t+i+20}, s_{t+i+42}, s_{t+i+60}, s_{t+i+79}, s_{t+i+94}, \cdots \right),$$

$$u_t = \left( \sum_{i \in A} s_{t+i}, \sum_{i \in A} s_{t+i}, \cdots, \sum_{i \in A} s_{t+i} \right),$$

$$y_t = \left( \sum_{i \in A} y_{t+i}, \sum_{i \in A} y_{t+i}, \cdots, \sum_{i \in A} y_{t+i} \right),$$

$$e_t = (e_t, e_{t+1}, \cdots, e_{t+m-1}).$$

Any even Hamming weight linear combination of Eq. (22) will generate a linear approximation without $\sum_{i \in A} s_{t+i}$ and $\sum_{i \in A} y_{t+i}$, which correlation would be treated as 0. As for odd linear combinations, it is still required that any of $\Lambda_i [1−3,5−8], i \in \mathbb{T}_i$ will not deduce a zero correlation for $h$ function. Therefore, we can construct a multidimensional linear approximation with dimension $9 \leq m \leq 42$, which consisting of $2^{n−1}−6 = 2^{m−7}$ linear approximations with correlation $2^{−57.045}$. By Theorem 1, its SEI $\Delta(e_t) = 2^{−121.0908}$.

As $s_t$ is a $m$-sequence, shifting and summation sequence $s'_{t+c_j} = s_{t+c_j} + \sum_{i \in A} s_{t+i}$ is also a $m$-sequence with same generator polynomial as $s_t$. Let vectorial sequence $x'_t = (s'_{t+c_1}, \cdots, s'_{t+c_m})$, since shift offsets $c_j, 1 \leq j \leq m$ have large difference, the parity-checks with $\tau = 1$ are not all ruled out.

Since $x'_t$ runs over $\mathbb{F}^m_2 \setminus \{0\}$, there is at most one parity-check with $\tau = 1$ for each $0 < n \leq N/m$. In order to increase the occurrence possibility for parity-check with $t = 1$, several redundant binary linear approximations with nonzero correlation could be added into the subspace. The dimension increases but SEI is almost unchanged. Therefore, the maximal probability value should decrease.

Another way is exploiting a kind of special parity-checks with $\tau > 1$. In order to avoid the great loss of SEI while implementing convolution, we play a trade-off trick when special parity-checks are feasible. For example, suppose we have $h$ special parity-checks as follows.

$$G_{n,1}x'_{t-d_n,1} + \sum_{i=1}^{a} G_{n-i,1}x'_{t-d_i} + Ee'_t = 0, \cdots, G_{n,h}x'_{t-d_n,h} + \sum_{i=1}^{a} G_{n-i,h}x'_{t-d_i} + Ee'_t = 0.$$

Notice that all of them involve vector variables $x'_t, x'_{t-d_1}, \cdots, x'_{t-d_n}$ except for the last variable $x'_{t-d_n}$. Let $D_{n-i,j} = G_{n-i,j} + G_{n-i,1}, 1 \leq i \leq a$, denote the coefficient difference between the $j$-th and the 1-st equation. Let $\sum_{i=1}^{a} D_{n-i,j}x'_{t-d_i} = \delta_j$ denote the difference value. Moreover, we require that $\delta_j$ satisfies some restrictions.

Since we have $h−1$ groups of linear equations with coefficients $(D_{n-1,j}, \cdots, D_{n-a,j})$, we require that those linear equation groups have the same solution subspace $S$ with large dimension, for example, with dimension $am−1$ or $am−2$, which implies that the rank of $(D_{n-1,j}, \cdots, D_{n-a,j})$ may be 1 or 2. Thus when $(x'_t, x'_{t-d_1}, \cdots, x'_{t-d_n}) \in S$, all $\delta_j = 0$. Otherwise, $\delta_j \neq 0$ are likely different. Thus we have

$$G_{n,1}x'_{t-d_n,1} + 0 + \sum_{i=1}^{a} G_{n-i,1}x'_{t-d_i} + Ee'_t = 0,$$

$$G_{n,2}x'_{t-d_n,2} + \delta_2 + \sum_{i=1}^{a} G_{n-i,2}x'_{t-d_i} + Ee'_t = 0,$$

$$\cdots,$$

$$G_{n,h}x'_{t-d_n,h} + \delta_h + \sum_{i=1}^{a} G_{n-i,h}x'_{t-d_i} + Ee'_t = 0.$$
Then the APP for $\sum_{i=1}^{a} G_{n-i,1} e_{t-d_i} + E e_t$ could be evaluated by total probability theorem according to whether all of $\delta_j$ are 0. The initial state is recovered from observed values $z_t$ of the error-corrected positions, i.e.,
\[
\sum_{i=1}^{a} G_{n-i,1} (x'_{t-d_i} + e_{t-d_i}) + E (x'_t + e_t) = \sum_{i=1}^{a} G_{n-i,1} x'_{t-d_i} + E x'_{t} = z_t.
\]

The dimension of linear approximation is not changed but the APP converges slower. Thus the decoding ability decreases when dimension of $S$ decreasing. However, the constraints for parity-checks is relaxed.

With these techniques, fast correlation attack could be performed with these special parity-checks and multidimensional linear approximations in (22).

### 5.3 Complexity Estimation

In this section, we estimate some theoretical bounds for Grain-128a, which would bring us a new perspective for its security margin.

Let the SEI $\Delta(e_t) = 2^{-\gamma'}$, dimension $m = 42$, and $p_0 = 2^{-m} + 2^{-\frac{m+1}{2}}$ be the maximal probability value. To simplify the process of estimating the expected number of positions with $p'_c > p'_0$, we need the following hypothesis.

**Hypothesis 1.**

- The probability distribution $P$ stemming from SEI is close to symmetric distribution $P$, i.e., $p_0$ is maximal and all other $p_i$ are nearly the same.
- There are at least 2 parity-checks with two taps, or there are more special parity-checks as stated in previous section.

Suppose we have $h$ special parity-checks corresponding to a solution subspace of dimension $am - 1$ as stated above. Let $v_1, \ldots, v_h$ denote the check values, $\gamma = (\gamma_0, \ldots, \gamma_{2^m-1})$ and $\gamma' = (\gamma'_0, \ldots, \gamma'_{2^m-1})$ denote the frequency values in $v_1, \ldots, v_h$ and $v_1, v_2 \oplus \delta_2, \ldots, v_h \oplus \delta_h$ respectively. There are two events that may deduce $p'_c > p'_0$: event $A$ denotes that $\gamma \in A'(\zeta)$, while event $B$ denotes that $\gamma' \in A'(\zeta)$. For simplicity, we only consider that when $A$ occurs, then we have

\[
M'_c = \frac{1}{2} p_c \left( \sum_{\gamma \in A'(\zeta)} \binom{h}{\gamma} \prod_{i} p_{i,\gamma} + \sum_{\gamma' \in A'(\zeta)} \binom{h}{\gamma'} \prod_{i} p'_{i,\gamma'} \right),
\]

\[
M'_0 = \frac{1}{2} p_0 \left( \sum_{\gamma \in A'(\zeta)} \binom{h}{\gamma} \prod_{i} p_{i,\gamma} + \sum_{\gamma' \in A'(\zeta)} \binom{h}{\gamma'} \prod_{i} p'_{i,\gamma'} \right).
\]

The first term denotes the probability that current noise symbol is $\zeta$ or 0, when the frequency vector $\gamma \in A'(\zeta)$ and all $\delta_j = 0$. The second term corresponds to when the frequency vector $\gamma \in A'(\zeta)$ but many $\delta_j \neq 0$, $2 \leq j \leq h$. Thus the observed vector is $\gamma'$. Since $\gamma'_i$ are likely different, it is reasonable to assume that the second terms of $M'_c$ and $M'_0$ are close. To simplify the evaluation, we only consider the first term.

Table 5 in Appendix A depicts the approximation of $M'_c(\frac{1}{2}$ is neglected). $M'_c$ is estimated by two methods: Skellam distribution and summation for small $l$. The two estimations are very close to each other. Let $D_l = M'_c - M'_0$ denote the difference. We also compute the summation $\sum_{l=1}^{36} M''_l$ and the difference summation $\sum_{l=1}^{36} D'_l$. For example, when $h = h_0 = 2$, the expected key stream length $N > 2^{48 + 42 + 1} = 2^{51}$. As $P$ is symmetric, it seems no need to evaluate every probability value of APP distribution. Therefore, we use the key stream length $N$ multiplying with the number of parity-checks $h$ as time complexity.

For the other case when there are at least 2 parity-checks with two taps, there is no probability loss caused by trade-off. The complexity estimation is similar.
6 Further Problems

The analysis of vectorial iterative decoding algorithm is very complicated, there are several problems needed further study.

Firstly, the time complexity is estimated by the key stream length multiplying with the number of parity-checks. There are lots of redundant computations. However, we have no idea whether FWHT acceleration technique could be applied in this case. Secondly, we don’t know the number of suitable parity-checks. Thus the estimation for $M'_i$ and $D'_i$ of Grain-128a is based on a hypothesis. Thirdly, in this paper, we didn’t study whether there is also K-tree like method to generate these parity-checks in matrix ring. Therefore, the complexity of the precomputation phase is skipped over.

7 Conclusion

In this paper, a vectorial iterative decoding algorithm for FCA is proposed. Two novel criterions are given to break tie and improve the decoding efficiency. The original binary FCA proposed by Meier and Staffelbach is a special case of our FCA with dimension 1. We describe some cryptographic properties about its statistical model, decoding efficiency etc. Based on the statistical property of the first iteration, we estimate the bound of expected key stream length from the perspective of iterative decoding. We also perform a scaled experiment to verify the validity of the vectorial iterative decoding algorithm.

Moreover, we apply it to stream cipher Grain-128a. We construct a multidimensional linear approximation with large SEI by bundling up those binary linear approximations proposed in CRYPTO 18. We also give an trade-off approach to use special parity-checks with more than 2 taps. Consequently, we give an estimation of data complexity for Grain-128a from the point view of vectorial iterative decoding, which is a novel result to evaluate the potential security margin of a real world cipher.

References


A Estimation of $M'_i$
Table 5: Estimation of some $M'_i$ with $m = 42$

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